

GENERATING FUNCTIONS AND CENTRAL LIMIT THEOREMS

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References

Standard Reference

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Sums of independent random variables and powers of generating functions

Coin tossing

- $\mathbb{P}\{ct = \text{head}\} = \mathbb{P}\{ct = \text{tail}\} = \frac{1}{2}$.
- random variable $\xi = \mathbb{I}_{\{ct=\text{tail}\}} = \begin{cases} 1 & \text{if tail} \\ 0 & \text{if head} \end{cases}$
- n independent runs: $\xi_1, \xi_2, \dots, \xi_n$, $\mathbb{P}\{\xi_j = 1\} = \mathbb{P}\{\xi_j = 0\} = \frac{1}{2}$.
- $X_n = \xi_1 + \xi_2 + \dots + \xi_n$... the number of tails within n runs

$$\mathbb{P}\{X_n = k\} = \frac{\binom{n}{k}}{2^n}$$

Sums of independent random variables and powers of generating functions

Counting generating function

$a_n = 2^n$... total number of possible n -runs

$a_{n,k} = \binom{n}{k}$... the number of n -runs with k tails

$A_n(u) = \sum_{k \geq 0} a_{n,k} u^k = \sum_{k \geq 0} \binom{n}{k} u^k = (1 + u)^n$... counting gen. func.

$A_n(1) = \sum_{k \geq 0} a_{n,k} = a_n = (1 + 1)^n = 2^n$

Sums of independent random variables and powers of generating functions

Probability generating function

$$\begin{aligned}\mathbb{E} u^{X_n} &= \sum_{k \geq 0} \mathbb{P}\{X_n = k\} \cdot u^k \\ &= \sum_{k \geq 0} \frac{1}{2^n} \binom{n}{k} \cdot u^k \\ &= \frac{(1+u)^n}{2^n} = \frac{A_n(u)}{A_n(1)}\end{aligned}$$

$$\mathbb{P}\{X_n = k\} = \frac{a_{n,k}}{a_n} \implies \boxed{\mathbb{E} u^{X_n} = \frac{A_n(u)}{A_n(1)}}$$

Sums of independent random variables and powers of generating functions

Powers of probability generating functions

$$\mathbb{E} u^\xi = \frac{1}{2} + \frac{1}{2}u = \frac{1+u}{2}$$

$$\begin{aligned} \implies \mathbb{E} u^{X_n} &= \mathbb{E} u^{\xi_1 + \dots + \xi_n} \\ &= \mathbb{E} (u^{\xi_1} \dots u^{\xi_n}) \\ &= \mathbb{E} (u^{\xi_1}) \dots \mathbb{E} (u^{\xi_n}) \quad \xi_j \text{ independent !!!} \\ &= \left(\frac{1+u}{2} \right)^n \end{aligned}$$

Sums of independent random variables and powers of generating functions

General fact

$X_n = \xi_1 + \xi_2 + \cdots + \xi_n$, where the r.v.'s ξ_j are **iid***

$$\implies \mathbb{E} u^{X_n} = \left(\mathbb{E} u^{\xi_1} \right)^n$$

* **Notation.** “iid” ... independently and identically distributed

Sums of independent random variables and powers of generating functions

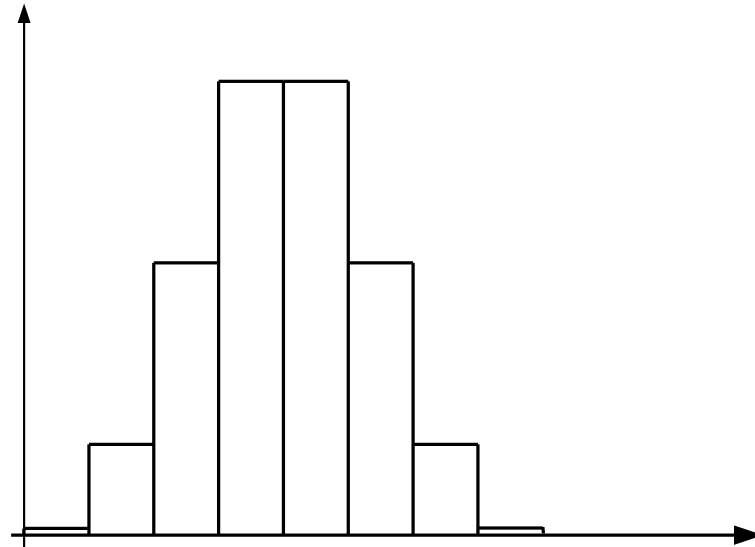
Relation to moment generating function $m_Z(v) = \mathbb{E} e^{vZ}$

$\mathbb{E}(Z^r)$... r -th moment of Z

$$\implies \sum_{r \geq 0} \mathbb{E}(Z^r) \frac{v^r}{r!} = \mathbb{E} \left(\sum_{r \geq 0} \frac{Z^r v^r}{r!} \right) = \mathbb{E} e^{vZ} = \mathbb{E} u^Z \quad \text{with } \boxed{u = e^v}.$$

A central limit theorem

Binomial coefficients

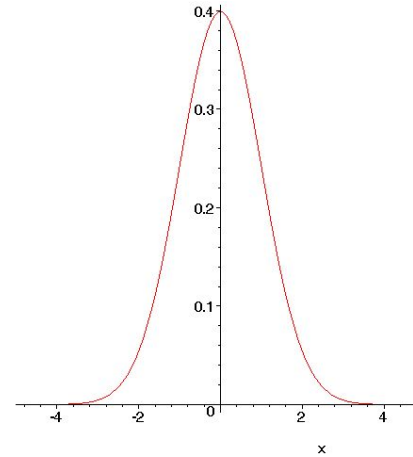


$$\binom{n}{k} = \frac{n!}{k!(n-k)!} = \frac{2^n}{\sqrt{\pi n/2}} \exp\left(-\frac{(k - \frac{n}{2})^2}{n/2}\right) + O(2^n/n)$$

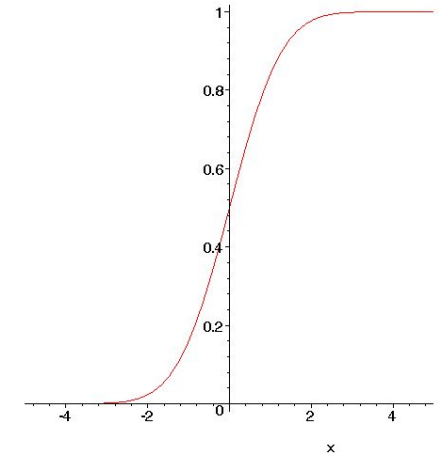
A central limit theorem

Standard normal distribution

density: $f(t) = \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}t^2}$.



normal distribution function: $\Phi(x) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^x e^{-\frac{1}{2}t^2} dt$



A central limit theorem

Normally distributed random variable

Definition

A random variable Z has **standard normal distribution** $N(0, 1)$ if

$$\mathbb{P}\{Z \leq x\} = \Phi(x).$$

A random variable Z is **normally distributed** (or **Gaussian**) with mean μ and variance σ^2 if its distribution function is given by

$$\mathbb{P}\{Z \leq x\} = \Phi\left(\frac{x - \mu}{\sigma}\right),$$

Notation. $\mathcal{L}(Z) = N(\mu, \sigma^2)$.

A central limit theorem

Moment generating function of $N(\mu, \sigma^2)$:

$$m_Z(v) = \mathbb{E} e^{vZ} = e^{\mu v - \frac{1}{2}\sigma^2 v^2}.$$

Characteristic function of $N(\mu, \sigma^2)$:

$$\varphi_Z(t) = \mathbb{E} e^{itZ} = e^{i\mu t - \frac{1}{2}\sigma^2 t^2}.$$

Standard normal distribution: $\mu = 0, \sigma^2 = 1$

$$\mathbb{E} e^{vZ} = e^{\frac{1}{2}v^2}, \quad \mathbb{E} e^{itZ} = e^{-\frac{1}{2}t^2}$$

A central limit theorem

Definition We say, that a sequence of random variables X_n satisfies a **central limit theorem** with (scaling) mean μ_n and (scaling) variance σ_n^2 if

$$\mathbb{P}\{X_n \leq \mu_n + x \cdot \sigma_n\} = \Phi(x) + o(1)$$

as $n \rightarrow \infty$.

Example. X_n = number of tails in n runs of coin tossing:

$$\begin{aligned} \mathbb{P}\{X_n \leq n/2 + x \cdot \sqrt{n/4}\} &= \sum_{k \leq n/2 + x \cdot \sqrt{n/4}} \frac{1}{2^n} \binom{n}{k} \\ &\sim \sum_{k \leq n/2 + x \cdot \sqrt{n/4}} \frac{1}{\sqrt{\pi n/2}} \exp\left(-\frac{(k - \frac{n}{2})^2}{n/2}\right) \sim \Phi(x). \end{aligned}$$

X_n satisfies a central limit theorem with mean $\frac{n}{2}$ and variance $\frac{n}{4}$.

Central Limit Theorem

Definition *Weak convergence:*

$$\boxed{X_n \xrightarrow{d} X} \quad :\Leftrightarrow \quad \boxed{\lim_{n \rightarrow \infty} \mathbb{P}\{X_n \leq x\} = \mathbb{P}\{X \leq x\}}$$

*for all points of continuity
of $F_X(x) = \mathbb{P}\{X \leq x\}$*

Reformulation:

X_n satisfies a **central limit theorem** with (scaling) mean μ_n and (scaling) variance σ_n^2 is the same as

$$\boxed{\frac{X_n - \mu_n}{\sigma_n} \xrightarrow{d} N(0, 1)}.$$

A central limit theorem

Weak convergence via moment generating functions

$$\lim_{n \rightarrow \infty} \mathbb{E} e^{vX_n} = \mathbb{E} e^{vX} \quad (v \in \mathbb{R}) \quad \implies \quad X_n \xrightarrow{d} X$$

Moreover, we have convergence of all moments: $\mathbb{E}(X_n^r) \rightarrow \mathbb{E}(X^r)$.

Recall: $\mathbb{E} e^{vX_n} = \mathbb{E}((e^v)^{X_n}) = \mathbb{E} u^{X_n}$ for $u = e^v$.

A central limit theorem

Weak convergence via characteristic functions (Levy's Criterion)

$$\lim_{n \rightarrow \infty} \mathbb{E} e^{itX_n} = \mathbb{E} e^{itX} \quad (t \in \mathbb{R}) \quad \iff \quad X_n \xrightarrow{d} X$$

Moreover, if for all $t \in \mathbb{R}$

$$\psi(t) := \lim_{n \rightarrow \infty} \mathbb{E} e^{itX_n}$$

exists and $\psi(t)$ is continuous at $t = 0$ then $\psi(t)$ is the characteristic function of a random variable X for which we have $X_n \xrightarrow{d} X$.

Central Limit Theorem

Theorem

ξ_1, ξ_2, \dots iid, $\mathbb{E} \xi_i^2 < \infty$, $X_n = \xi_1 + \xi_2 + \dots + \xi_n$

$$\implies \boxed{\frac{X_n - \mathbb{E} X_n}{\sqrt{\mathbb{V} X_n}} \xrightarrow{d} N(0, 1)}$$

Remark. $\iff \mathbb{P}\{X_n \leq \mathbb{E} X_n + x\sqrt{\mathbb{V} X_n}\} = \Phi(x) + o(1)$.

Proof

$$\mu = \mathbb{E} \xi_i, \sigma^2 = \mathbb{V} \xi_i = \mathbb{E} (\xi_i^2) - (\mathbb{E} \xi_i)^2 \implies \mathbb{E} X_n = n\mu, \mathbb{V} X_n = n\sigma^2.$$

Central Limit Theorem

$$\varphi_{\xi_i}(t) = \mathbb{E} e^{it\xi_i} = e^{it\mu - \frac{1}{2}\sigma^2 t^2} (1+o(1)) \quad (t \rightarrow 0)$$

$$\varphi_{X_n}(t) = \varphi_{\xi_i}(t)^n$$

$$Z_n := (X_n - \mu n) / \sqrt{\sigma^2 n}$$

$$\begin{aligned} \implies \boxed{\varphi_{Z_n}(t)} &= \mathbb{E} e^{itZ_n} \\ &= e^{-it\sqrt{n}\mu/\sigma} \cdot \mathbb{E} \left(e^{(it/(\sqrt{n}\sigma))(\xi_1 + \dots + \xi_n)} \right) \\ &= e^{-it\sqrt{n}\mu/\sigma} \cdot \boxed{\left(\mathbb{E} e^{(it/(\sqrt{n}\sigma)\xi_1)} \right)^n} \\ &= e^{-it\sqrt{n}\mu/\sigma} \cdot e^{it\sqrt{n}\mu/\sigma - \frac{1}{2}t^2} (1+o(1)) \\ &= e^{-\frac{1}{2}t^2} (1+o(1)) \rightarrow \boxed{e^{-\frac{1}{2}t^2}}. \end{aligned}$$

+ Levy's criterion.

A central limit theorem

Quasi-Power Theorem (Hwang)

Let X_n be a sequence of random variables with the property that

$$\mathbb{E} u^{X_n} = A(u) \cdot B(u)^{\lambda_n} \cdot \left(1 + O\left(\frac{1}{\phi_n}\right) \right)$$

holds uniformly in a complex neighborhood of $u = 1$, $\lambda_n \rightarrow \infty$ and $\phi_n \rightarrow \infty$, and $A(u)$ and $B(u)$ are analytic functions in a neighborhood of $u = 1$ with $A(1) = B(1) = 1$. Set

$$\mu = B'(1) \quad \text{and} \quad \sigma^2 = B''(1) + B'(1) - B'(1)^2.$$

$$\implies \mathbb{E} X_n = \mu \lambda_n + O(1 + \lambda_n/\phi_n), \quad \mathbb{V} X_n = \sigma^2 \lambda_n + O(1 + \lambda_n/\phi_n),$$

$$\frac{X_n - \mathbb{E} X_n}{\sqrt{\mathbb{V} X_n}} \xrightarrow{d} N(0, 1) \quad (\sigma^2 \neq 0).$$

Bivariate generating functions

Bivariate counting generating function

$$A(x, u) = \sum_{n, k \geq 0} \binom{n}{k} u^k x^n = \sum_{n \geq 0} (1 + u)^n x^n = \frac{1}{1 - x(1 + u)}.$$

Observation: this is a **rational function!**

Bivariate generating functions

Rational functions

$P(x, u), Q(x, u)$ polynomials:

$$A(x, u) = \sum_{n,k \geq 0} a_{n,k} u^k x^n = \frac{P(x, u)}{Q(x, u)}$$

Assumption: factorization of denominator

$$Q(x, u) = \prod_{j=1}^r \left(1 - \frac{x}{\rho_j(u)} \right)$$

with

$$\boxed{|\rho_1(u)| < \max_{2 \leq j \leq r} |\rho_j(u)|} \quad \text{for } |u - 1| < \varepsilon.$$

Bivariate generating functions

Central limit theorem for rational functions

Suppose that $A(x, u) = \sum a_{n,k} u^k x^n$ with $a_{n,k} \geq 0$ is **rational** and satisfies the assumptions from above.

Let X_n be a sequence of random variables with

$$\mathbb{P}\{X_n = k\} = \frac{a_{n,k}}{a_n}$$

with $a_n = \sum_k a_{n,k}$.

Then X_n satisfies a **central limit theorem** with

$$\mu_n = -n \frac{\rho_1'(1)}{\rho_1(1)} \quad \text{and} \quad \sigma_n^2 = n \left(-\frac{\rho_1''(1)}{\rho_1(1)} - \frac{\rho_1'(1)}{\rho_1(1)} + \frac{\rho_1'(1)^2}{\rho_1(1)^2} \right).$$

Bivariate generating functions

Proof

Partial fraction decomposition:

$$A(x, u) = \frac{C_1(u)}{1 - x/\rho_1(u)} + \dots + \frac{C_r(u)}{1 - x/\rho_r(u)}$$

$$\implies A_n(u) = \sum_{k \geq 0} a_{n,k} u^k = C_1(u) \rho_1(u)^{-n} + \dots + C_r(u) \rho_r(u)^{-n} \sim C_1(u) \rho_1(u)^{-n}$$

$$\implies \mathbb{E} u^{X_n} = \frac{A_n(u)}{A_n(1)} \sim \frac{C_1(u)}{C_1(1)} \left(\frac{\rho_1(1)}{\rho_1(u)} \right)^n$$

\implies central limit theorem.

Bivariate generating functions

Integer compositions

$3 = 1 + 1 + 1 = 2 + 1 = 1 + 2 = 3 \dots$ 4 compositions of 3.

$a_n =$ number of compositions of n , $A(x) = \sum a_n x^n$:

$$A(x) = 1 + A(x)(x + x^2 + x^3 + \dots) = 1 + A(x)\frac{x}{1-x}.$$

$$\implies A(x) = \frac{1}{1 - \frac{x}{1-x}} = \frac{1-x}{1-2x}$$

$$\implies \boxed{a_n = 2^{n-1}}$$

Bivariate generating functions

Integer compositions

$a_{n,k}$ = number of integer composition of n with k summands

$$A(x, u) = \sum a_{n,k} u^k x^n:$$

$$A(x, u) = 1 + uA(x, u)(x + x^2 + x^3 + \dots) = 1 + A(x, u) \frac{xu}{1-x}.$$

$$\implies A(x, u) = \frac{1}{1 - \frac{xu}{1-x}} = \frac{1-x}{1-x(1+u)}$$

\implies **central limit theorem** with $\mu_n = \frac{n}{2}$ and $\sigma^2 = \frac{n}{4}$.

Bivariate generating functions

Systems of linear equations

Suppose, that several generating functions

$$A_1(x, u) = \sum_{n,k} a_{1;n,k} u^k x^n, \dots, A_r(x, u) = \sum_{n,k} a_{r;n,k} u^k x^n$$

satisfy a **linear system of equations**.

Then all generating functions $A_j(x, u)$ are rational and a **central limit theorem** for corresponding random variables is **expected**.

Bivariate generating functions

Meromorphic functions

The function $A(x, u)$ is meromorphic in x when u is considered as a parameter and there exists a dominant root $\rho_1(u)$ such that (locally)

$$A(x, u) = \frac{C(x, u)}{1 - \frac{x}{\rho_1(u)}}$$

$$\implies A_n(u) \sim C(\rho_1(u), u) \cdot \rho_1(u)^{-n}$$

$$\implies \mathbb{E} u^{X_n} \sim \frac{C(\rho_1(u), u)}{C(\rho_1(1), 1)} \left(\frac{\rho_1(1)}{\rho_1(u)} \right)^n$$

\implies central limit theorem.

Bivariate generating functions

Number of cycles in permutations

$p_{n,k}$ = number of permutations of $\{1, 2, \dots, n\}$ with k cycles

$$\hat{P}(x, u) = \sum_{n, k \geq 0} p_{n,k} \cdot u^k \cdot \frac{x^n}{n!} = e^{u \cdot \log \frac{1}{1-x}} = \frac{1}{(1-x)^u}$$

Remark: $p_{n,k} = (-1)^{n-k} s_{n,k}$, where $s_{n,k}$ are the **Stirling number of the first kind**.

Excursion: Singularity Analysis

Lemma 1 *Suppose that*

$$y(x) = (1 - x/x_0)^{-\alpha}.$$

Then

$$y_n = (-1)^n \binom{-\alpha}{n} x_0^{-n} = \frac{n^{\alpha-1}}{\Gamma(\alpha)} x_0^{-n} + \mathcal{O}(n^{\alpha-2}) x_0^{-n}.$$

Remark: This asymptotic expansion is uniform in α if α varies in a compact region of the complex plane.

Excursion: Singularity Analysis

Lemma 2 (Flajolet and Odlyzko) *Let*

$$y(x) = \sum_{n \geq 0} y_n x^n$$

be analytic in a region

$$\Delta = \{x : |x| < x_0 + \eta, |\arg(x - x_0)| > \delta\},$$

$$x_0 > 0, \eta > 0, 0 < \delta < \pi/2.$$

Suppose that for some real α

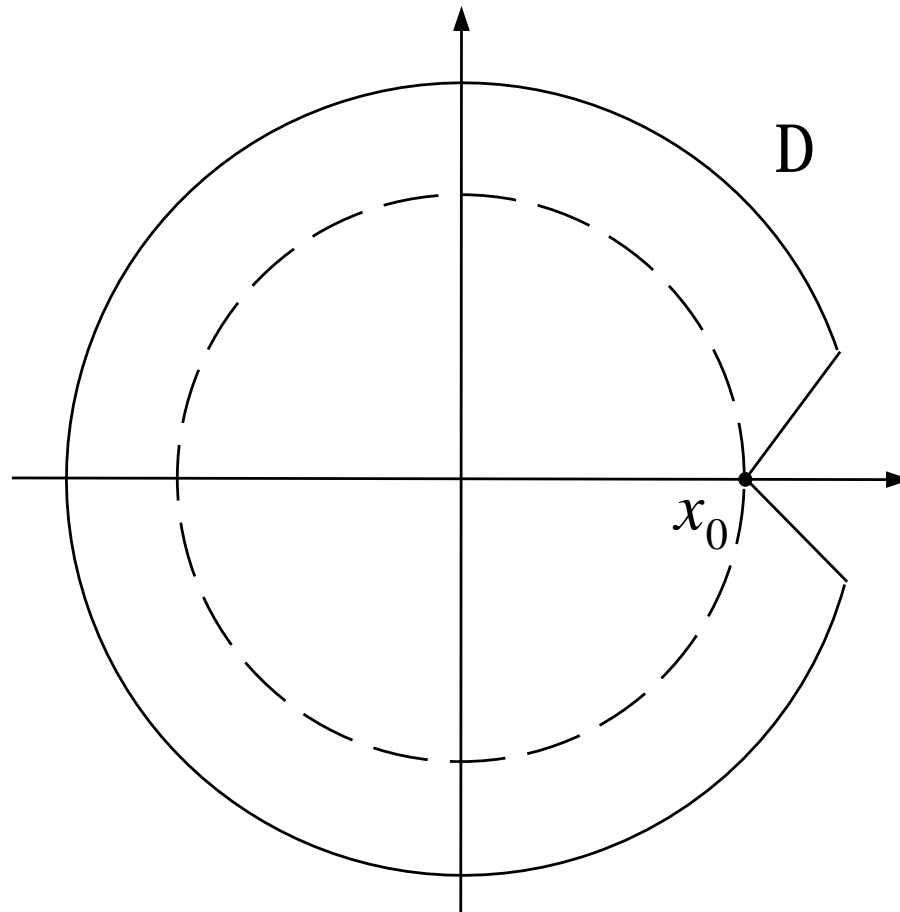
$$y(x) = \mathcal{O}\left((1 - x/x_0)^{-\alpha}\right) \quad (x \in \Delta).$$

Then

$$y_n = \mathcal{O}\left(x_0^{-n} n^{\alpha-1}\right).$$

Excursion: Singularity Analysis

Δ -region



Bivariate generating functions

Number of cycles in permutations (continued)

$$\hat{P}(x, u) = e^{u \log \frac{1}{1-x}} = \frac{1}{(1-x)^u}$$

$$\begin{aligned} \implies p_n(u) &= \sum_{k \geq 0} p_{n,k} u^k \\ &\sim n! \frac{n^{u-1}}{\Gamma(u)} \\ &= n! \frac{e^{(u-1) \log n}}{\Gamma(u)} \end{aligned}$$

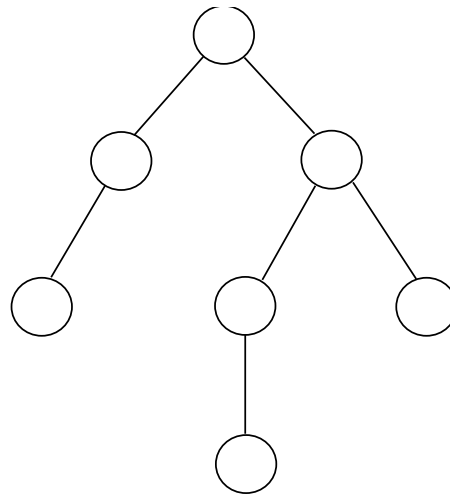
$$\implies \boxed{\mathbb{E} u^{X_n} \sim \frac{1}{\Gamma(u)} (e^{u-1})^{\log n}}$$

\implies central limit theorem with $\mu_n = \log n$ and $\sigma_n^2 = \log n$.

Generalization: Exp-Log-Schemes: $F(x, u) = e^{h(u) \log \frac{1}{1-x} + R(x, u)}$.

Bivariate generating functions

Catalan trees g_n = number of Catalan trees of size n .



$$G(x) = x(1 + G(x) + G(x)^2 + \dots) = \frac{x}{1 - G(x)}$$

$$G(x) = \frac{1 - \sqrt{1 - 4x}}{2} \implies g_n = \frac{1}{n} \binom{2n - 2}{n - 1}.$$

(Catalan numbers)

Bivariate generating functions

Catalan trees with singularity analysis

$$G(x) = \frac{1 - \sqrt{1 - 4x}}{2} = \frac{1}{2} - \frac{1}{2}\sqrt{1 - 4x}$$

$$\implies g_n \sim -\frac{1}{2} \cdot \frac{4^n n^{-3/2}}{\Gamma(-\frac{1}{2})} = \frac{4^{n-1}}{\sqrt{\pi} \cdot n^{3/2}}$$

Bivariate generating functions

Number of leaves of Catalan trees

$g_{n,k}$ = number of Catalan trees of size n with k leaves.

$$G(x, u) = xu + x(G(x, u) + G(x, u)^2 + \dots) = xu + \frac{xG(x, u)}{1 - G(x, u)}$$

$$\implies G(x, u) = \frac{1}{2} \left(1 + (u - 1)x - \sqrt{1 - 2(u + 1)x + (u - 1)^2 x^2} \right)$$

$$\implies G(x, u) = g(x, u) - h(x, u) \sqrt{1 - \frac{x}{\rho(u)}}$$

for certain analytic function $g(x, u)$, $h(x, u)$, and $\rho(u)$.

Bivariate generating functions

Application of singularity analysis

Considering u as a parameter we get

$$G_n(u) = \sum_{k \geq 0} g_{n,k} u^k \sim \frac{h(\rho(u), u) \cdot \rho(u)^{-n} \cdot n^{-3/2}}{2\sqrt{\pi}}$$

$$\implies \boxed{\mathbb{E} u^{X_n} = \frac{G_n(u)}{G_n(1)} \sim \frac{h(\rho(u), u)}{h(\rho(1), 1)} \left(\frac{\rho(1)}{\rho(u)} \right)^n}$$

\implies central limit theorem with $\mu_n = \frac{n}{2}$ and $\sigma_n^2 = \frac{n}{8}$

Bivariate generating functions

Cayley trees

$r_{n,k}$ = number of Cayley trees of size n with k leaves

$$R(x, u) = \sum_{n,k \geq 0} r_{n,k} u^k \frac{x^n}{n!}$$

$$\implies \boxed{R(x, u) = x e^{R(x, u)} + x(u - 1)}$$

$$\implies \text{?????}$$

Functional equations

Catalan trees: $G(x, u) = xu + xG(x, u)/(1 - G(x, u))$

Cayley trees: $R(x, u) = xe^{R(x, u)} + x(u - 1)$

Recursive structure leads to functional equation for gen. func.:

$$A(x, u) = \Phi(x, u, A(x, u))$$

Functional equations

Linear functional equation: $\Phi(x, u, a) = \Phi_0(x, u) + a\Phi_1(x, u)$

$$\implies A(x, u) = \frac{\Phi_0(x, u)}{1 - \Phi_1(x, u)}$$

Usually techniques similar to those used for rational resp. meromorphic functions work and prove a **central limit theorem**.

Functional equations

Non-linear functional equations: $\Phi_{aa}(x, u, a) \neq 0$.

Suppose that $A(x, u) = \Phi(x, u, A(x, u))$, where $\Phi(x, u, a)$ has a power series expansion at $(0, 0, 0)$ with non-negative coefficients and $\Phi_{aa}(x, u, a) \neq 0$.

Let $x_0 > 0$, $a_0 > 0$ (inside the region of convergence) satisfy the system of equations:

$$a_0 = \Phi(x_0, 1, a_0), \quad 1 = \Phi_a(x_0, 1, a_0).$$

Then there exists analytic function $g(x, u)$, $h(x, u)$, and $\rho(u)$ such that locally

$$A(x, u) = g(x, u) - h(x, u) \sqrt{1 - \frac{x}{\rho(u)}}.$$

Functional equations

Idea of the Proof.

Set $F(x, u, a) = \Phi(x, u, a) - a$. Then we have

$$F(x_0, 1, a_0) = 0$$

$$F_a(x_0, 1, a_0) = 0$$

$$F_x(x_0, 1, a_0) \neq 0$$

$$F_{aa}(x_0, 1, a_0) \neq 0.$$

Weierstrass preparation theorem implies that there exist analytic functions $H(x, u, a)$, $p(x, u)$, $q(x, u)$ with $H(x_0, 1, a_0) \neq 0$, $p(x_0, 1) = q(x_0, 1) = 0$ and

$$F(x, u, a) = H(x, u, a) \left((a - a_0)^2 + p(x, u)(a - a_0) + q(x, u) \right).$$

Functional equations

$$F(x, u, a) = 0 \iff (a - a_0)^2 + p(x, u)(a - a_0) + q(x, u) = 0.$$

Consequently

$$\begin{aligned} A(x, u) &= a_0 - \frac{p(x, u)}{2} \pm \sqrt{\frac{p(x, u)^2}{4} - q(x, u)} \\ &= \boxed{g(x, u) - h(x, u) \sqrt{1 - \frac{x}{\rho(u)}}}, \end{aligned}$$

where we write

$$\frac{p(x, u)^2}{4} - q(x, u) = K(x, u)(x - \rho(u))$$

which is again granted by the Weierstrass preparation theorem and we set

$$g(x, u) = a_0 - \frac{p(x, u)}{2} \quad \text{and} \quad h(x, u) = \sqrt{-K(x, u)\rho(u)}.$$

Functional equations

A central limit theorem for functional equations

Suppose that $A(x, u) = \Phi(x, u, A(x, u))$, where $\Phi(x, u, a)$ has a power series expansion at $(0, 0, 0)$ with non-negative coefficients and $\Phi_{aa}(x, u, a) \neq 0$ (+ *minor* technical conditions). Set

$$\mu = \frac{x_0 \Phi_x(x_0, 1, a_0)}{\Phi(x_0, 1, a_0)} \quad \text{and} \quad \sigma^2 = \text{"long formula"}.$$

Then the random variable X_n defined by $\mathbb{P}\{X_n = k\} = a_{n,k}/a_n$ satisfies a **central limit theorem** with

$$\mu_n = n\mu \quad \text{and} \quad \sigma_n^2 = n\sigma^2.$$

Functional equations

Number of leaves in Cayley trees ($R(x) = xe^{T(x)}$)

$$R(x, u) = xe^{R(x, u)} + x(u - 1)$$

$$x_0 = \frac{1}{e}, \quad r_0 = R(x_0) = 1.$$

\implies central limit theorem with $\mu_n = \frac{1}{e}n$ and $\sigma^2 = \frac{e-2}{e^2}n$.

Functional equations

Systems of functional equations

Suppose, that several generating functions

$$A_1(x, u) = \sum_{n,k} a_{1;n,k} u^k x^n, \dots, A_r(x, u) = \sum_{n,k} a_{r;n,k} u^k x^n$$

satisfy a **system of non-linear equations**

$$A_j(x, u) = \Phi_j(x, u, A_1(x, u), \dots, A_r(x, u)),$$

where $\Phi_j(x, u, a_1, \dots, a_r)$ is non-linear in a_1, \dots, a_r for some j and has a power series expansion at $(0, 0, 0)$ with non-negative coefficients (for all j).

Let $x_0 > 0$, $\mathbf{a}_0 = (a_{0,0}, \dots, a_{r,0}) > 0$ (inside the region of convergence) satisfy the system of equations: $(\Phi = (\Phi_1, \dots, \Phi_r))$

$$\mathbf{a}_0 = \Phi(x_0, 1, \mathbf{a}_0), \quad 0 = \det(\mathbb{I} - \Phi_{\mathbf{a}}(x_0, 1, \mathbf{a}_0)).$$

Functional equations

Suppose further, that the **dependency graph** of the system $\mathbf{a} = \Phi(x, u, \mathbf{a})$ is **strongly connected**.

Then there exists analytic function $g_j(x, u)$, $h_j(x, u)$, and $\rho(u)$ (that is **independent of j**) such that locally

$$A_j(x, u) = g_j(x, u) - h_j(x, u) \sqrt{1 - \frac{x}{\rho(u)}}.$$

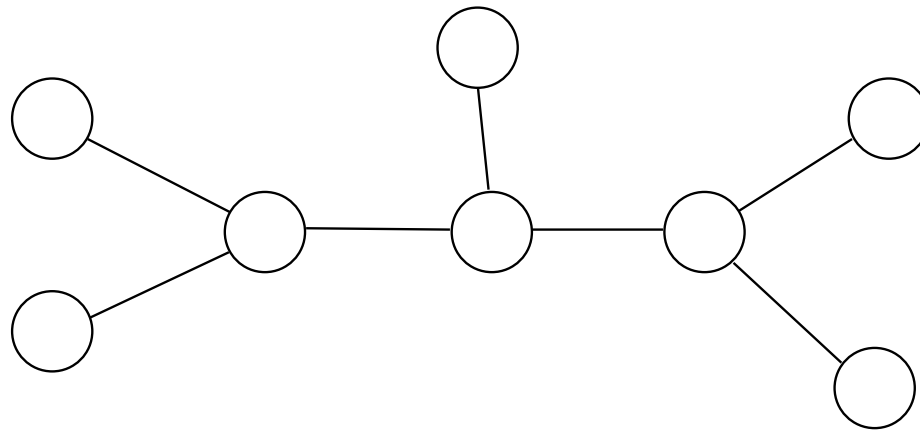
If $A(x, u) = \sum_{n,k} a_{n,k} x^n u^k = F(x, u, A_1(x, u), \dots, A_j(x, u))$ (for some analytic function F satisfying certain conditions) then then random variable X_n defined by $\mathbb{P}\{X_n = k\} = a_{n,k}/a_n$ satisfies a **central limit theorem** with

$$\mu_n = n\mu \quad \text{and} \quad \sigma_n^2 = n\sigma^2,$$

where μ and σ^2 can be computed.

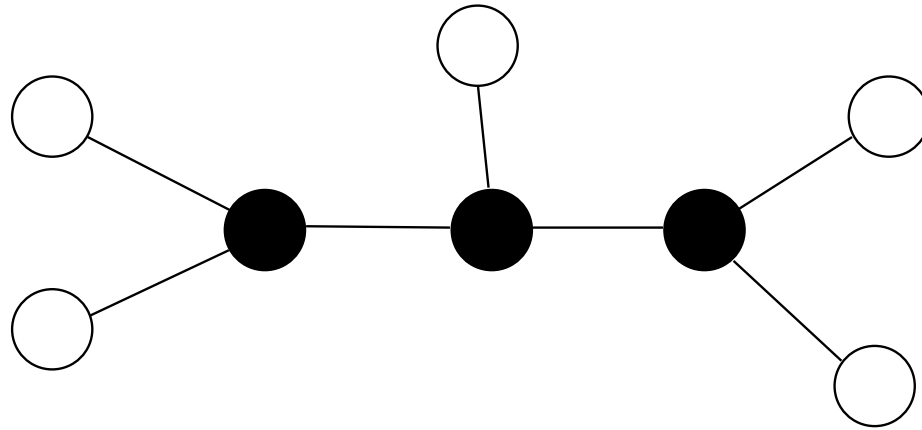
Patterns in Trees

Pattern \mathcal{M}



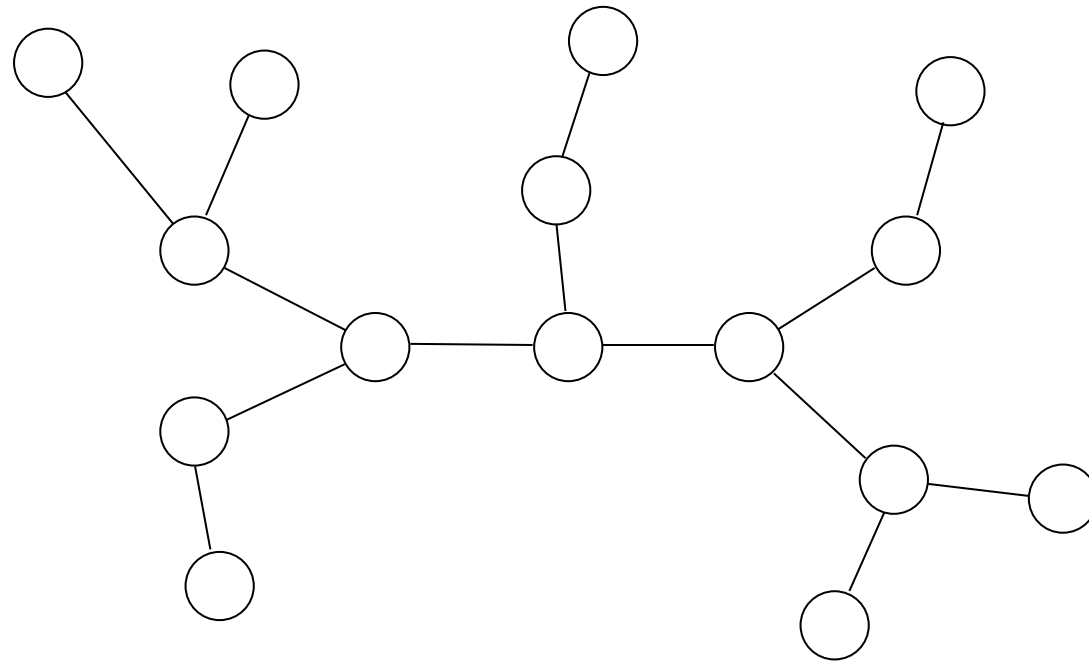
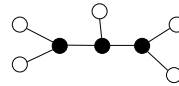
Patterns in Trees

Pattern \mathcal{M}



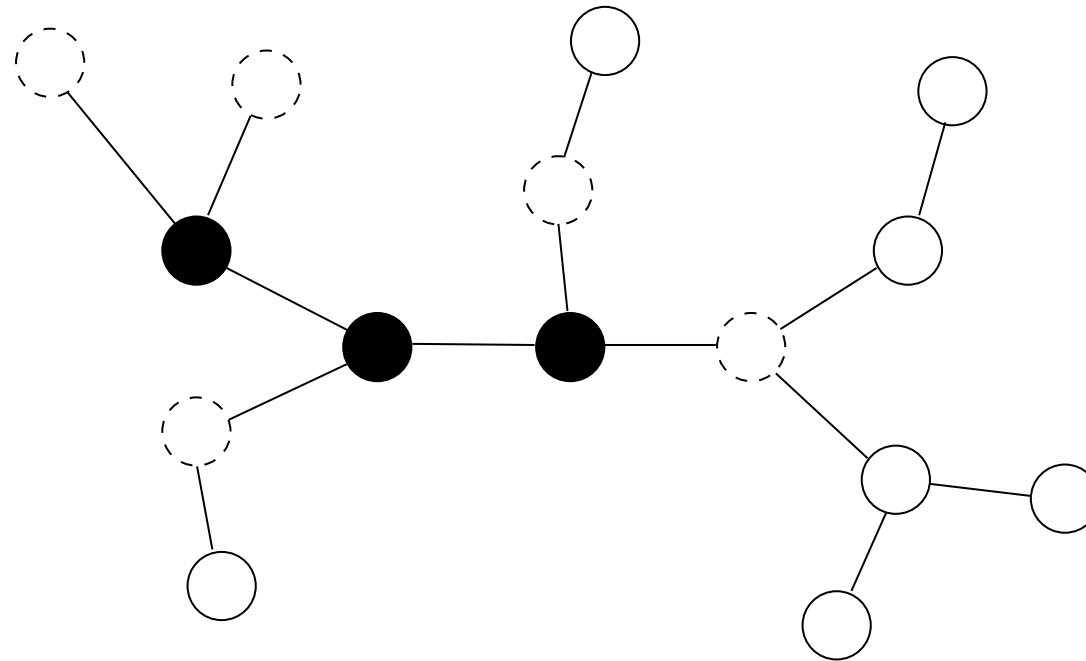
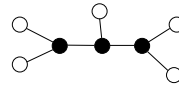
Patterns in Trees

Occurrence of a pattern \mathcal{M}



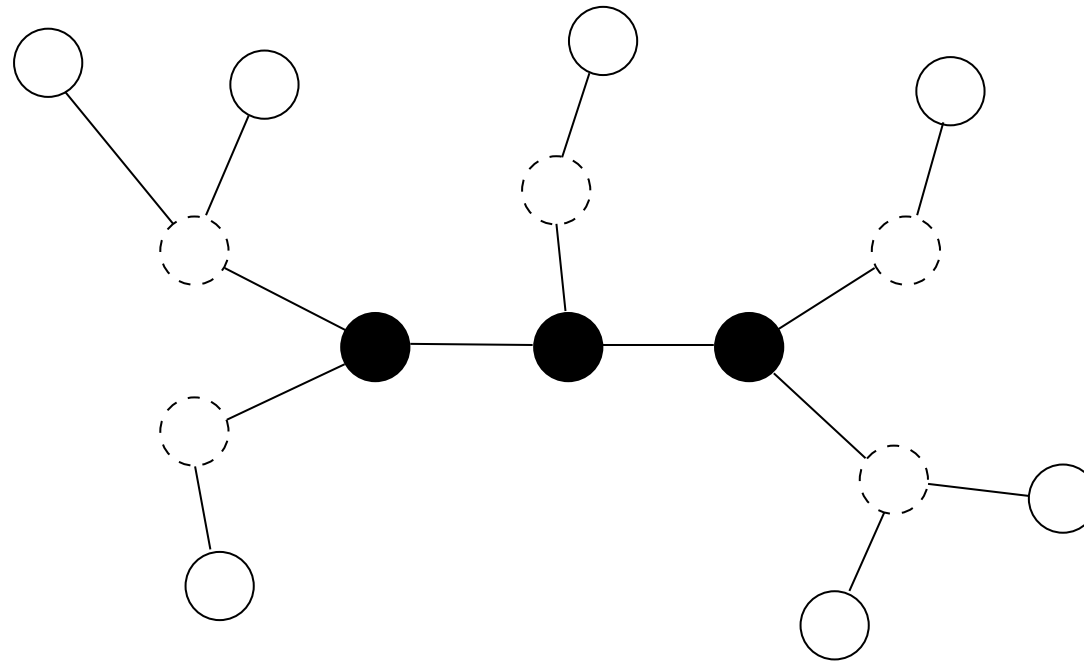
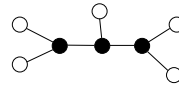
Patterns in Trees

Occurrence of a pattern \mathcal{M}



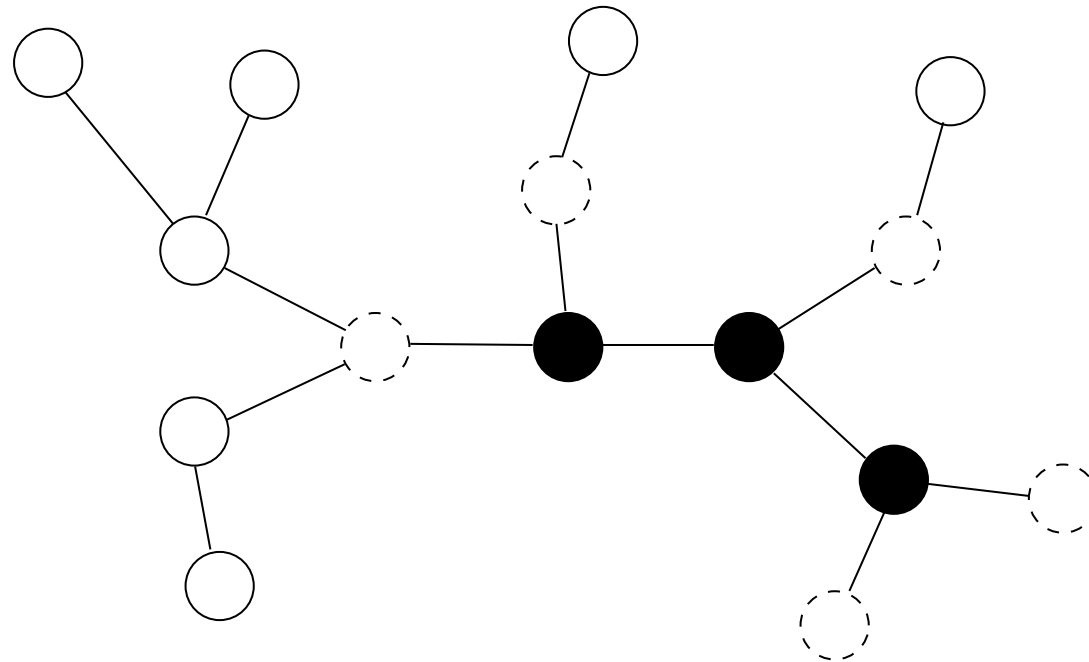
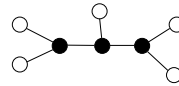
Patterns in Trees

Occurrence of a pattern \mathcal{M}



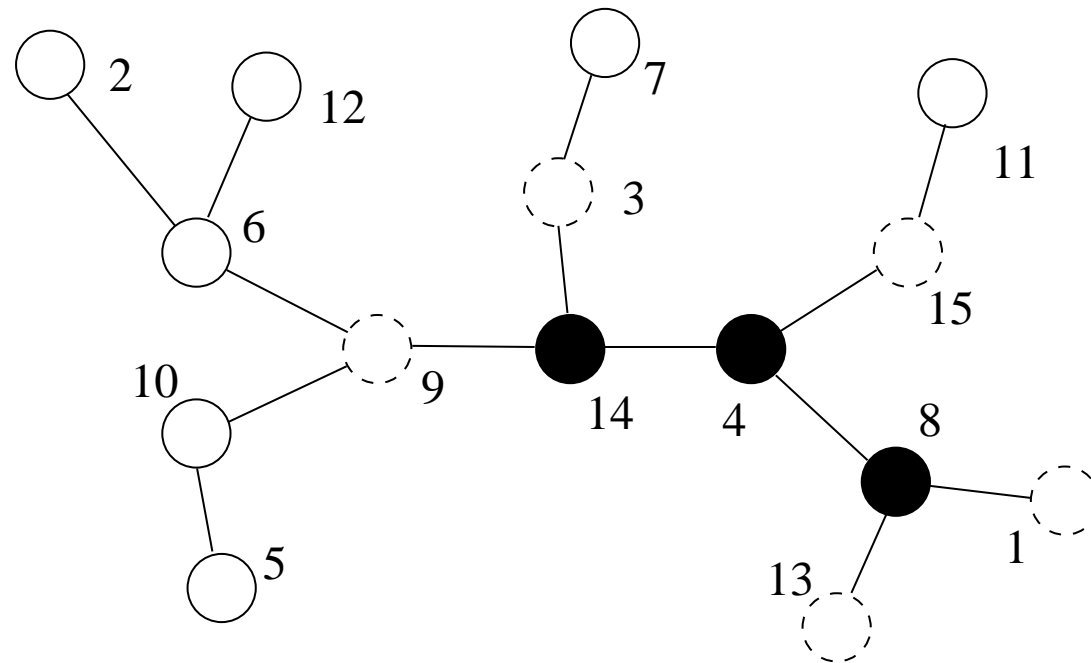
Patterns in Trees

Occurrence of a pattern \mathcal{M}



Patterns in Trees

Occurrence of a pattern \mathcal{M}  in a labelled tree



Patterns in Trees

Cayley's formula

$r_n = n^{n-1}$... number of **rooted** labelled trees with n nodes

$t_n = n^{n-2}$... number of labelled trees with n nodes

Generating functions

$$R(x) = \sum_{n \geq 1} r_n \frac{x^n}{n!}:$$

$$R(x) = xe^{R(x)}$$

$$t(x) = \sum_{n \geq 1} t_n \frac{x^n}{n!}:$$

$$T(x) = R(x) - \frac{1}{2}R(x)^2$$

Patterns in Trees

Theorem

\mathcal{M} ... be a given finite tree.

X_n ... number of occurrences of \mathcal{M} in a labelled tree of size n

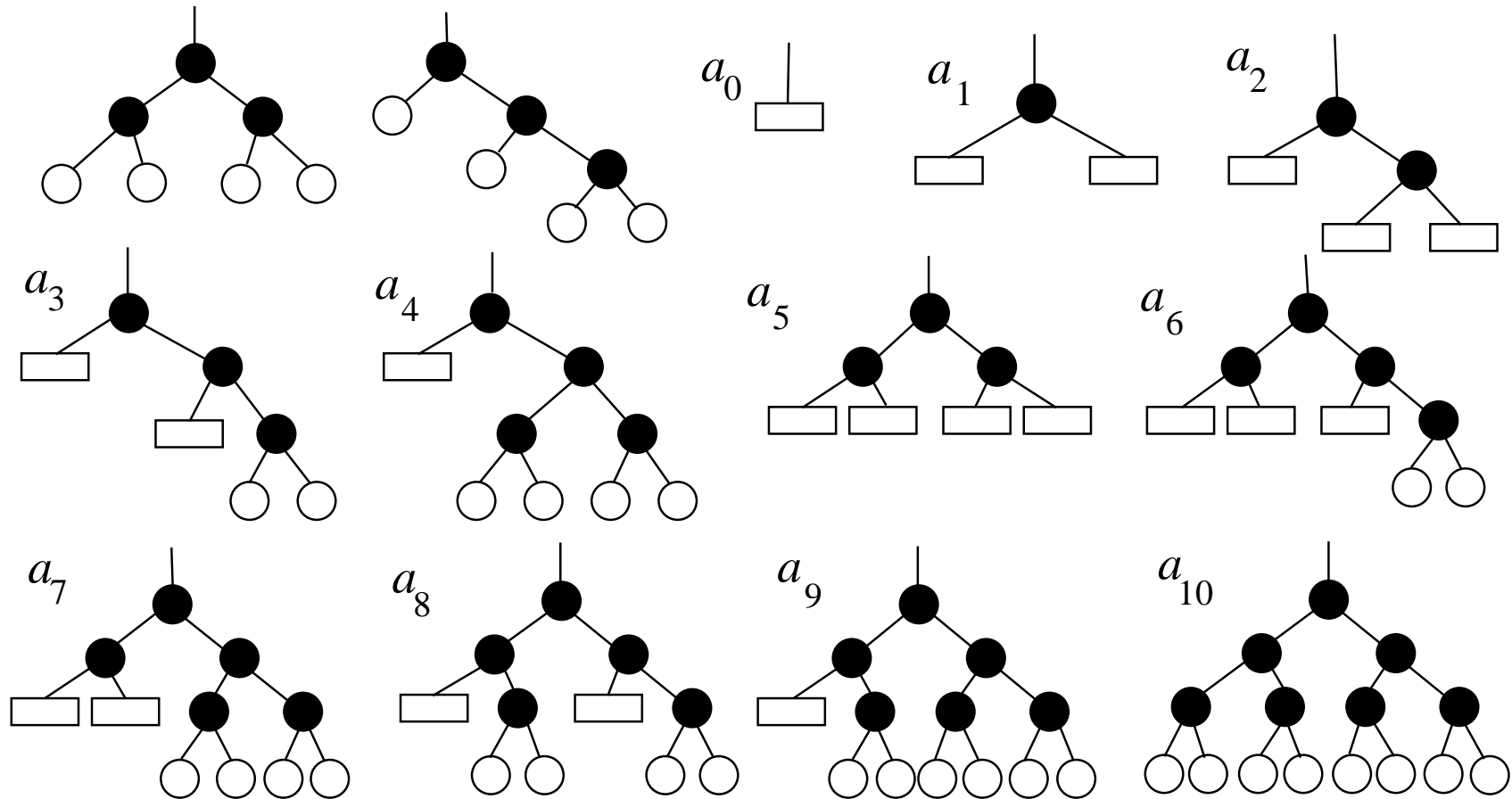
$\implies X_n$ satisfies a **central limit theorem** with

$$\mathbb{E} X_n \sim \mu n \quad \text{and} \quad \mathbb{V} X_n \sim \sigma^2 n.$$

$\mu > 0$ and $\sigma^2 \geq 0$ depend on the pattern \mathcal{M} and can be computed explicitly and algorithmically and can be represented as polynomials (with rational coefficients) in $1/e$.

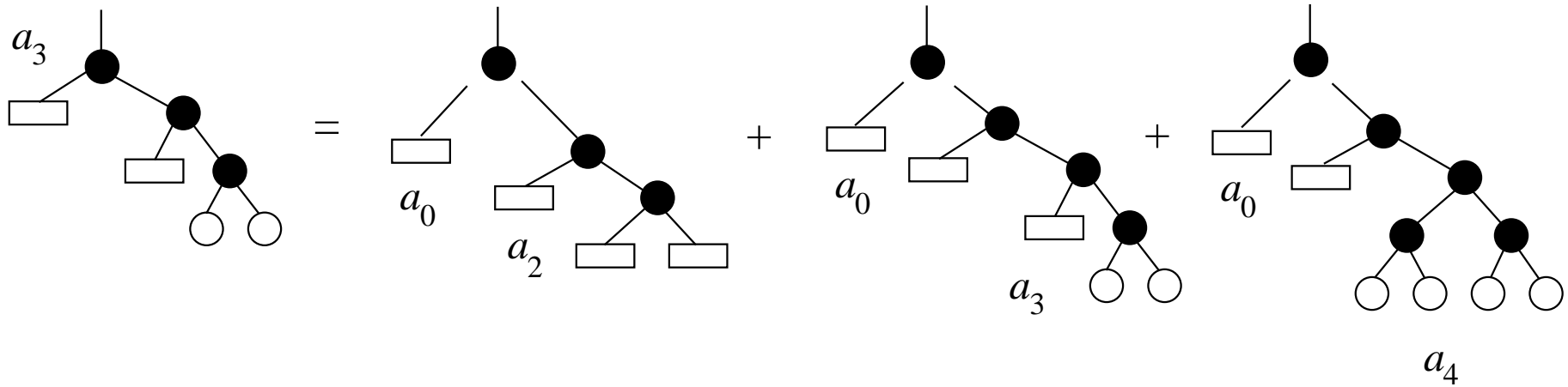
Patterns in Trees

Partition of trees in classes (\square ... out-degree different from 2)



Patterns in Trees

Recurrences $A_3 = xA_0A_2 + xA_0A_3 + xA_0A_4$

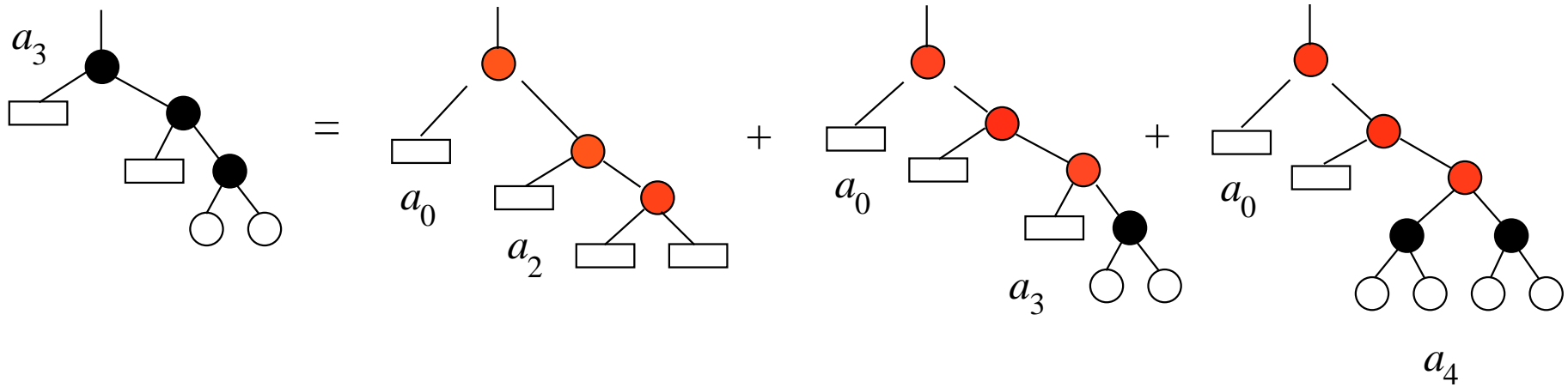


$$A_j(x) = \sum_{n,k} a_{j;n} \frac{x^n}{n!}$$

$a_{j;n}$... number of trees of size n in class j

Patterns in Trees

Recurrences $A_3 = xuA_0A_2 + xuA_0A_3 + xuA_0A_4$



$$A_j(x, u) = \sum_{n,k} a_{j;n,k} \frac{x^n}{n!} u^k$$

$a_{j;n,k}$... number of trees of size n in class j with k occurrences of \mathcal{M}

Patterns in Trees

$$A_0 = A_0(x, u) = x + x \sum_{i=0}^{10} A_i + x \sum_{n=3}^{\infty} \frac{1}{n!} \left(\sum_{i=0}^{10} A_i \right)^n,$$

$$A_1 = A_1(x, u) = \frac{1}{2}x A_0^2,$$

$$A_2 = A_2(x, u) = x A_0 A_1,$$

$$A_3 = A_3(x, u) = x A_0 (A_2 + A_3 + A_4)u,$$

$$A_4 = A_4(x, u) = x A_0 (A_5 + A_6 + A_7 + A_8 + A_9 + A_{10})u^2,$$

$$A_5 = A_5(x, u) = \frac{1}{2}x A_1^2 u,$$

$$A_6 = A_6(x, u) = x A_1 (A_2 + A_3 + A_4)u^2,$$

$$A_7 = A_7(x, u) = x A_1 (A_5 + A_6 + A_7 + A_8 + A_9 + A_{10})u^3,$$

$$A_8 = A_8(x, u) = \frac{1}{2}x (A_2 + A_3 + A_4)^2 u^3,$$

$$A_9 = A_9(x, u) = x (A_2 + A_3 + A_4) (A_5 + A_6 + A_7 + A_8 + A_9 + A_{10})u^4,$$

$$A_{10} = A_{10}(x, u) = \frac{1}{2}x (A_5 + A_6 + A_7 + A_8 + A_9 + A_{10})^2 u^5.$$

Patterns in Trees

Final Result for $\mathcal{M} =$ 

Central limit theorem with

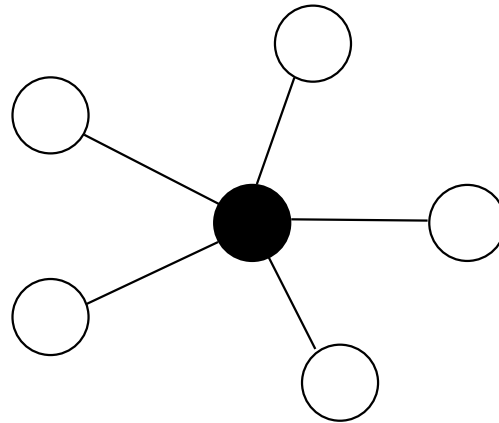
$$\mu = \frac{5}{8e^3} = 0.0311169177\dots$$

and

$$\sigma^2 = \frac{20e^3 + 72e^2 + 84e - 175}{32e^6} = 0.0764585401\dots$$

Degree distribution in random trees

Nodes of given degree



$$l = 5$$

$X_n^{(l)}$ = number of nodes of degree l in trees of size n satisfies a **central limit theorem** with

$$\mu^{(l)} = \frac{1}{e(l-1)!} \quad \text{and} \quad (\sigma^{(l)})^2 = \frac{1 + (l-2)^2}{e^2(l-1)!^2} + \frac{1}{e(l-1)!}$$

Degree distribution in random trees

$d_{n,\ell}$... probability that a random node in a random labelled tree of size n has degree ℓ :

$$\mathbb{E} X_n^{(\ell)} = n d_{n,\ell}$$

$$d_\ell := \lim_{n \rightarrow \infty} d_{n,\ell} = \frac{1}{e(\ell-1)!} = \mu^{(\ell)}$$

Probability generating function of the degree distribution:

$$p(w) := \sum_{\ell \geq 1} d_\ell w^\ell = w e^{w-1}$$

Degree distribution in random planar graphs

- **Outerplanar graph:** no K_4 and no $K_{2,3}$ as a minor.
- **Series-parallel graph:** no K_4 as a minor.
- **Planar graph:** no K_5 and no $K_{3,3}$ as a minor.

Remark.

outerplanar \subseteq series-parallel \subseteq planar

Degree distribution in random planar graphs

Outerplanar Graphs

g_n ... number of **labelled outer-planar** graphs with n vertices:

$$G(x) = \sum_{n \geq 0} g_n \frac{x^n}{n!}$$

$$\begin{aligned} G(x) &= e^{C(x)}, \\ C'(x) &= e^{B'(xC'(x))}, \\ B'(x) &= x + \frac{1}{2}x A(x), \\ A(x) &= x(1 + A(x))^2 + x(1 + A(x))A(x) \\ &= \frac{1 - 3x - \sqrt{1 - 6x + x^2}}{4x}. \end{aligned}$$

Degree distribution in random planar graphs

Series-Parallel Graphs

$g_{n,m}$... number of **labelled series-parallel** graphs with n vertices and m edges:

$$G(x, y) = \sum_{n \geq 0} g_{n,m} \frac{x^n}{n!} y^m$$

$$G(x, y) = e^{C(x,y)}$$

$$\frac{\partial C(x, y)}{\partial x} = \exp \left(\frac{\partial B}{\partial x} \left(x \frac{\partial C(x, y)}{\partial x}, y \right) \right),$$

$$\frac{\partial B(x, y)}{\partial y} = \frac{x^2}{2} \frac{1 + D(x, y)}{1 + y},$$

$$D(x, y) = (1 + y)e^{S(x,y)} - 1,$$

$$S(x, y) = (D(x, y) - S(x, y))xD(x, y).$$

Degree distribution in random planar graphs

Planar Graphs

$g_{n,m}$... number of **labelled planar** graphs with n vertices and m edges:

$$G(x, y) = \sum_{n \geq 0} g_{n,m} \frac{x^n}{n!} y^m$$

$$\begin{aligned} G(x, y) &= \exp(C(x, y)), \\ \frac{\partial C(x, y)}{\partial x} &= \exp\left(\frac{\partial B}{\partial x}\left(x \frac{\partial C(x, y)}{\partial x}, y\right)\right), \\ \frac{\partial B(x, y)}{\partial y} &= \frac{x^2}{2} \frac{1 + D(x, y)}{1 + y}, \\ \frac{M(x, D)}{2x^2 D} &= \log\left(\frac{1 + D}{1 + y}\right) - \frac{x D^2}{1 + x D}, \\ M(x, y) &= x^2 y^2 \left(\frac{1}{1 + xy} + \frac{1}{1 + y} - 1 - \frac{(1 + U)^2 (1 + V)^2}{(1 + U + V)^3} \right), \\ U &= xy(1 + V)^2, \\ V &= y(1 + U)^2. \end{aligned}$$

Degree distribution in random planar graphs

[Gimenez+Noy (2005)]

g_n number of all labelled planar graphs

$$g_n \sim c \cdot n^{-\frac{7}{2}} \gamma^n n!, \quad \gamma = 27.22\dots$$

Degree distribution in random planar graphs

Outerplanar graphs

Theorem

$X_n^{(\ell)}$... number of vertices of degree ℓ in random 2-connected, connected or unrestricted **labelled outerplanar** graphs with n vertices.

$\implies X_n^{(\ell)}$ satisfies a **central limit theorem** with

$$\mathbb{E} X_n^{(\ell)} \sim \mu^{(\ell)} n \quad \text{and} \quad \mathbb{V} X_n^{(\ell)} \sim (\sigma^{(\ell)})^2 n.$$

Degree distribution in random planar graphs

Outerplanar graphs $d_\ell = \mu^{(\ell)}$, $p(w) = \sum_{\ell \geq 1} d_\ell w^\ell$

- **2-connected**

$$p(w) = \frac{2(3 - 2\sqrt{2})w^2}{(1 - (\sqrt{2} - 1)w)^2}$$

- **connected or unrestricted:**

$$p(w) = \frac{c_1 w^2}{(1 - c_2 w)^2} \exp\left(c_3 w + \frac{c_4 w^2}{(1 - c_2 w)}\right)$$

(with certain constants $c_1, c_2, c_3, c_4 > 0$).

Degree distribution in random planar graphs

Series-parallel graphs

Theorem

$X_n^{(\ell)}$... number of vertices of degree ℓ in random 2-connected, connected or unrestricted **labelled series-parallel** graphs with n vertices.

$\implies X_n^{(\ell)}$ satisfies a **central limit theorem** with

$$\mathbb{E} X_n^{(\ell)} \sim \mu^{(\ell)} n \quad \text{and} \quad \mathbb{V} X_n^{(\ell)} \sim (\sigma^{(\ell)})^2 n.$$

Degree distribution in random planar graphs

2-connected series-parallel graphs $d_\ell = \mu^{(\ell)}$, $p(w) = \sum_{\ell \geq 1} d_\ell w^\ell$:

$$p(w) = \frac{B_1(1, w)}{B_1(1, 1)},$$

where $B_1(y, w)$ is given by the following procedure ...

Degree distribution in random planar graphs

$$\frac{E_0(y)^3}{E_0(y) - 1} = \left(\log \frac{1 + E_0(y)}{1 + R(y)} - E_0(y) \right)^2,$$

$$R(y) = \frac{\sqrt{1 - 1/E_0(y)} - 1}{E_0(y)},$$

$$E_1(y) = - \left(\frac{2R(y)E_0(y)^2(1 + R(y)E_0(y))^2}{(2R(y)E_0(y) + R(y)^2E_0(y)^2)^2 + 2R(y)(1 + R(y)E_0(y))} \right)^{\frac{1}{2}},$$

$$D_0(y, w) = (1 + yw)e^{\frac{R(y)E_0(y)}{1+R(y)E_0(y)}D_0(y,w)} - 1,$$

$$D_1(y, w) = \frac{(1 + D_0(y, w)) \frac{R(y)E_1(y)D_0(y,w)}{1+R(y)E_0(y)}}{1 - (1 + D_0(y, w)) \frac{R(y)E_0(y)D_0(y,w)}{1+R(y)E_0(y)}},$$

$$B_0(y, w) = \frac{R(y)D_0(y, w)}{1 + R(y)E_0(y)} - \frac{R(y)^2E_0(y)D_0(y, w)^2}{2(1 + R(y)E_0(y))},$$

$$B_1(y, w) = \frac{R(y)D_1(y, w)}{1 + R(y)E_0(y)} - \frac{R(y)^2E_0(y)D_0(y, w)D_1(y, w)}{1 + R(y)E_0(y)} - \frac{R(y)^2E_1(y)D_0(y, w)(1 + D_0(y, w)/2)}{(1 + R(y)E_0(y))^2}.$$

Degree distribution in random planar graphs

Theorem

$X_n^{(\ell)}$... number of vertices of degree ℓ in random 3-connected, 2-connected, connected or unrestricted **labelled planar** graphs with n vertices.

$$\implies \mathbb{E} X_n^{(\ell)} \sim \mu^{(\ell)} n$$

For $\ell \leq 2$, $X_n^{(\ell)}$ satisfies also a central limit theorem.

Degree distribution in random planar graphs

unrestricted planar graphs $d_\ell = \mu^{(\ell)}$, $p(w) = \sum_{\ell \geq 1} d_\ell w^\ell$:

d_1	d_2	d_3	d_4	d_5	d_6
0.0367284	0.1625794	0.2354360	0.1867737	0.1295023	0.0861805

$$p(w) = -e^{B_0(1,w)-B_0(1,1)} B_2(1,w) + e^{B_0(1,w)-B_0(1,1)} \frac{1 + B_2(1,1)}{B_3(1,1)} B_3(1,w)$$

where $B_j(y, w)$ are given by the following procedure ...

Degree Distribution

- Implicit equation for $D_0(y, w)$:

$$1 + D_0 = (1 + y \boxed{w}) \exp \left(\frac{\sqrt{S}(D_0(t-1) + t)}{4(3t+1)(D_0+1)} - \frac{D_0^2(t^4 - 12t^2 + 20t - 9) + D_0(2t^4 + 6t^3 - 6t^2 + 10t - 12) + t^4 + 6t^3 + 9t^2}{4(t+3)(D_0+1)(3t+1)} \right),$$

where $t = t(y)$ satisfies $y+1 = \frac{1+2t}{(1+3t)(1-t)} \exp \left(-\frac{1}{2} \frac{t^2(1-t)(18+36t+5t^2)}{(3+t)(1+2t)(1+3t)^2} \right)$.
and $S = (D_0(t-1) + t)(D_0(t-1)^3 + t(t+3)^2)$.

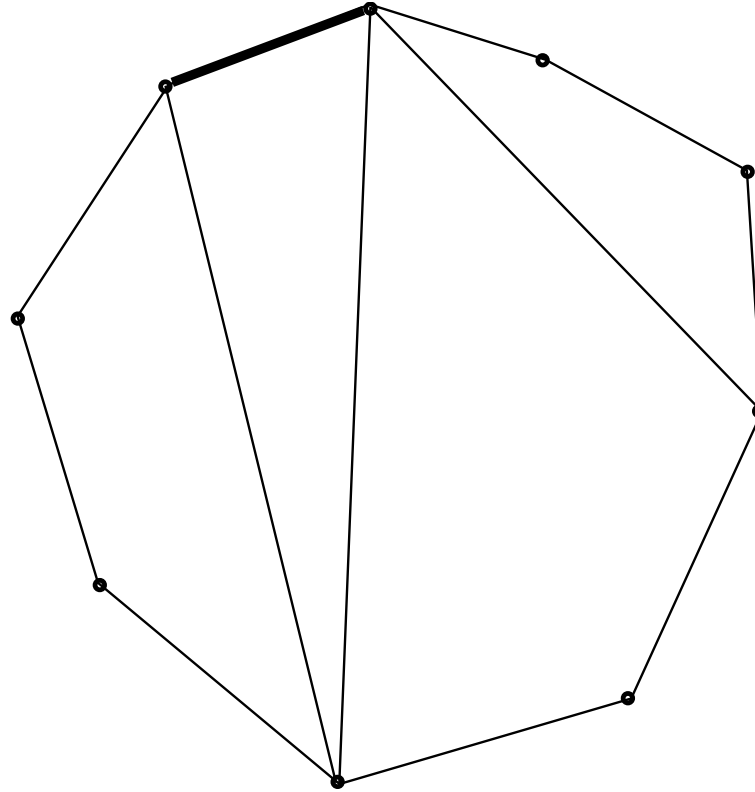
- Explicit expressions in terms of $D_0(y, w)$:

$$D_2(y, w), D_3(y, w), B_0(y, w), B_2(y, w), B_3(y, w)$$

- Explicit expression for $p(w)$:

$$p(w) = -e^{B_0(1,w)-B_0(1,1)} B_2(1, w) + e^{B_0(1,w)-B_0(1,1)} \frac{1 + B_2(1, 1)}{B_3(1, 1)} B_3(1, w)$$

Dissections



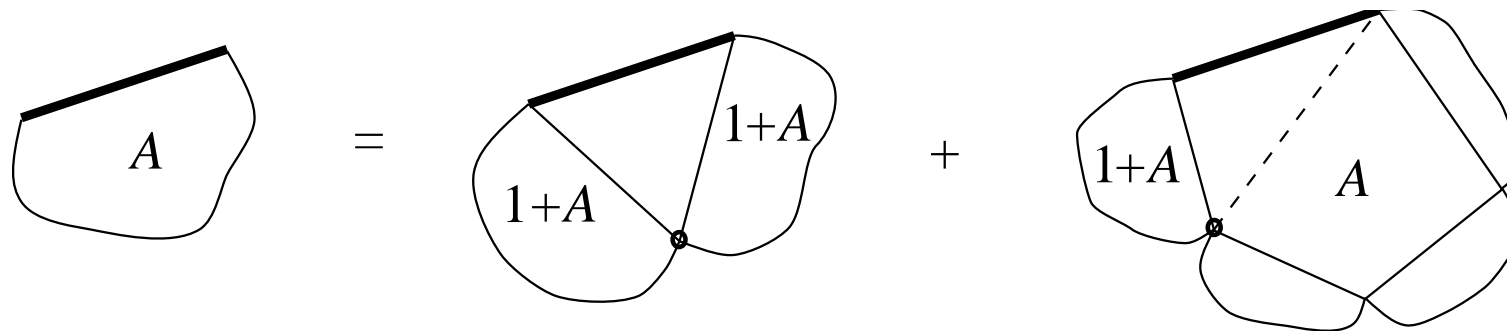
\mathcal{A} ... set of dissections

(unlabelled planar graphs, where all nodes are on the outer face, one edge is marked, and there are at least 3 edges)

Dissections

a_n ... number of dissections with $n + 2$ nodes, $n \geq 1$,
(the nodes of the marked edge are not counted)

$A(x) = \sum_{n \geq 1} a_n x^n$... generating function of dissections



$$A(x) = x(1 + A(x))^2 + x(1 + A(x))A(x)$$

Dissections

Quadratic equation:

$$A^2 + \frac{3x-1}{2x}A + \frac{1}{2} = 0$$

Solution:

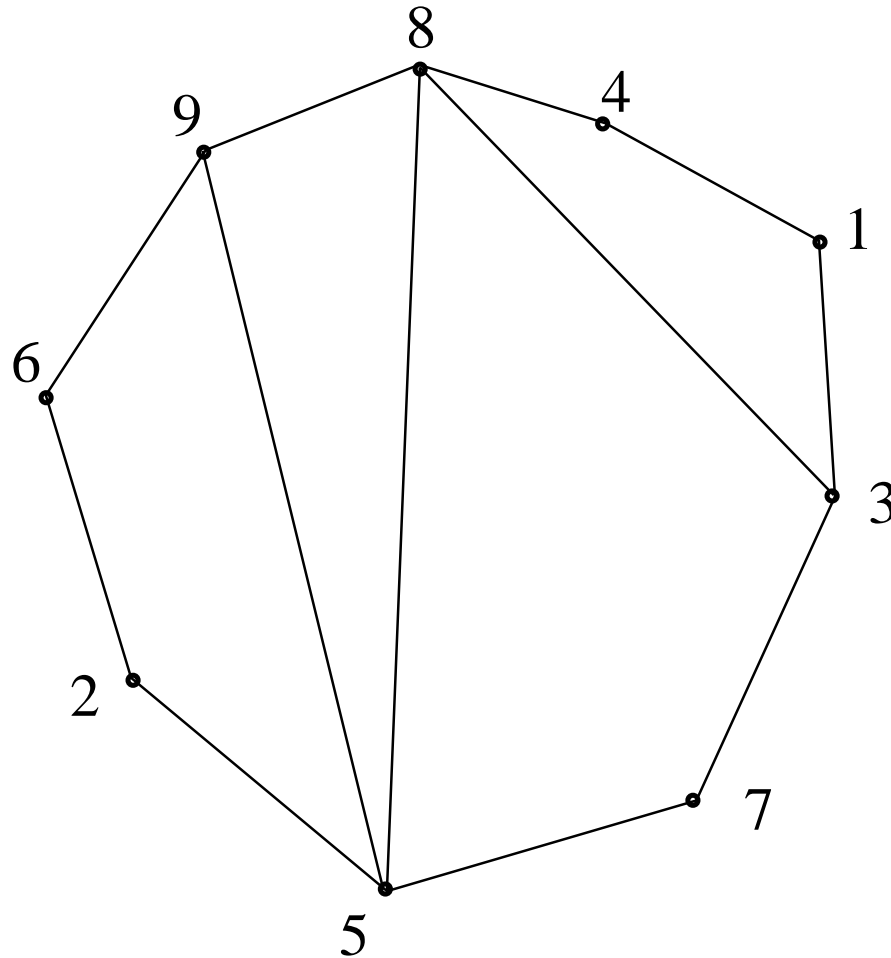
$$A(x) = \frac{1 - 3x - \sqrt{1 - 6x + x^2}}{4x}$$

Radius of convergence: $\rho_1 = 3 - 2\sqrt{2}$.

Lagrange inversion formula:

$$a_n = \frac{1}{n} \sum_{\ell=0}^{n-1} \binom{n}{\ell} \binom{n}{\ell+1} 2^\ell.$$

2-Connected Outer Planar Graphs



b_n ... number of 2-connected vertex labelled outer planar graphs

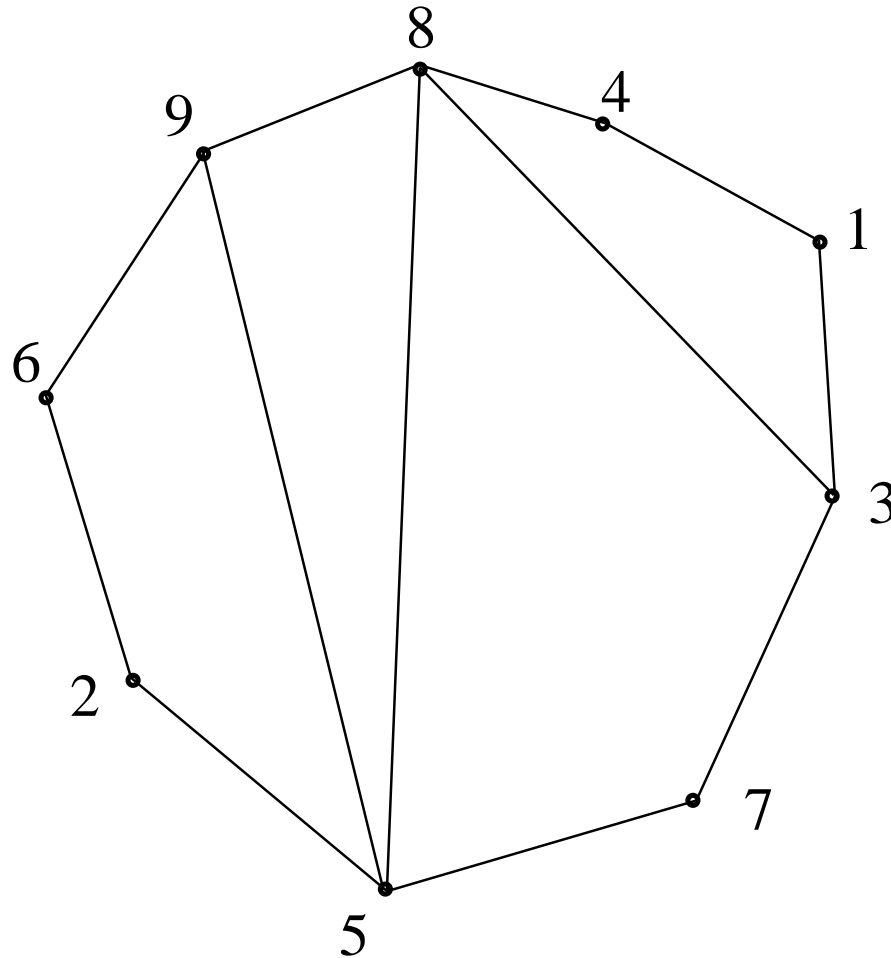
2-Connected Outer Planar Graphs

$B(x) = \sum_{n \geq 1} b_n \frac{x^n}{n!}$... exponential generating function of 2-connected labelled outer planar graphs:

$$B'(x) = x + \frac{1}{2}xA(x)$$

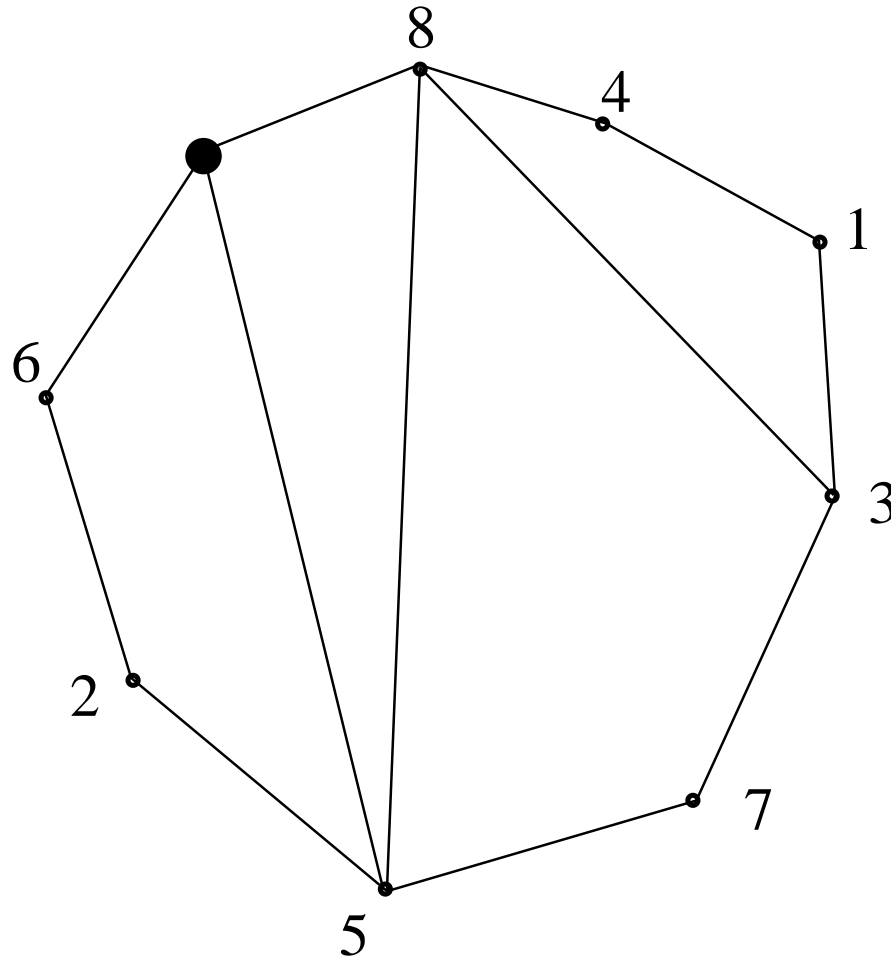
The derivative $B'(x)$ can be also interpreted as the exponential generating function $B^\bullet(x)$ of 2-connected labelled outer planar graphs, where one node is marked (and is not counted).

2-Connected Outer Planar Graphs



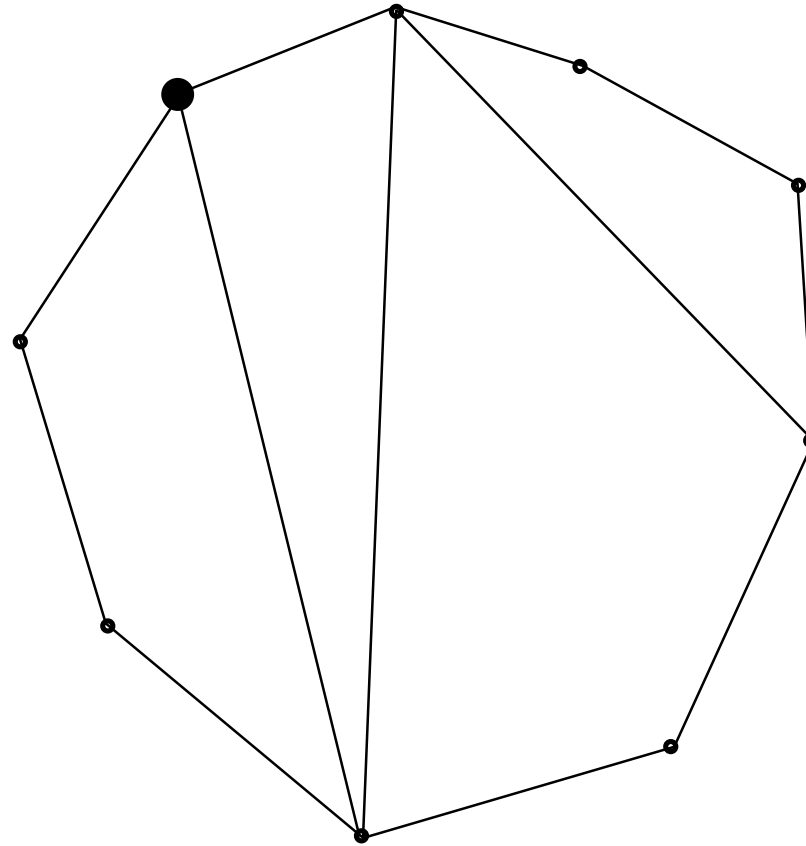
$$b_n = \frac{1}{2} a_{n-2} \cdot (n-1)! \quad (n \geq 3)$$

2-Connected Outer Planar Graphs



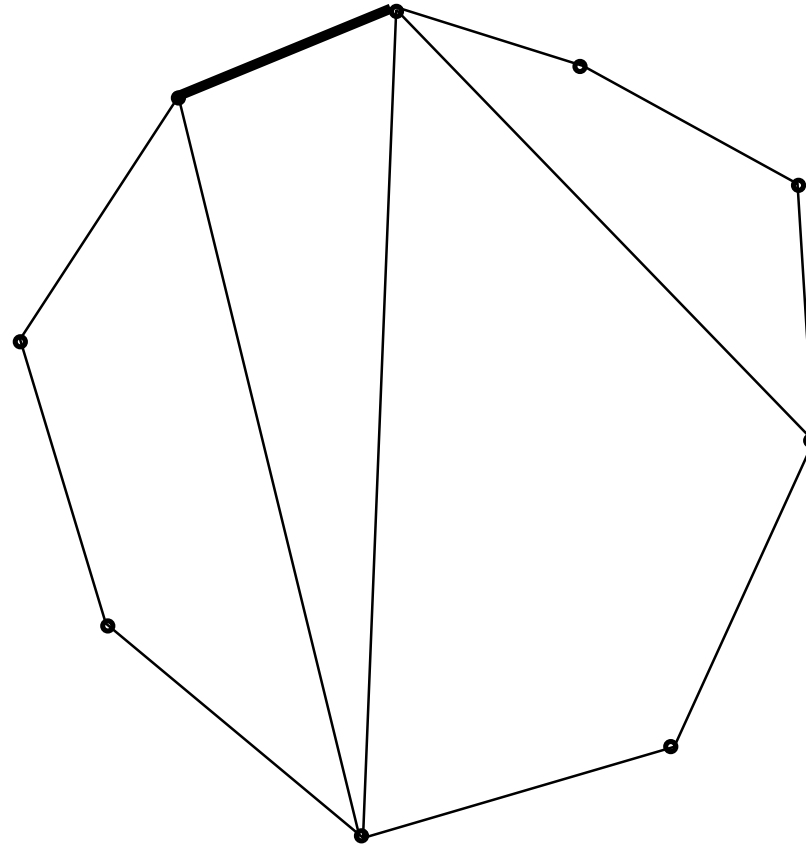
$$b_n = \frac{1}{2} a_{n-2} \cdot (n-1)! \quad (n \geq 3)$$

2-Connected Outer Planar Graphs



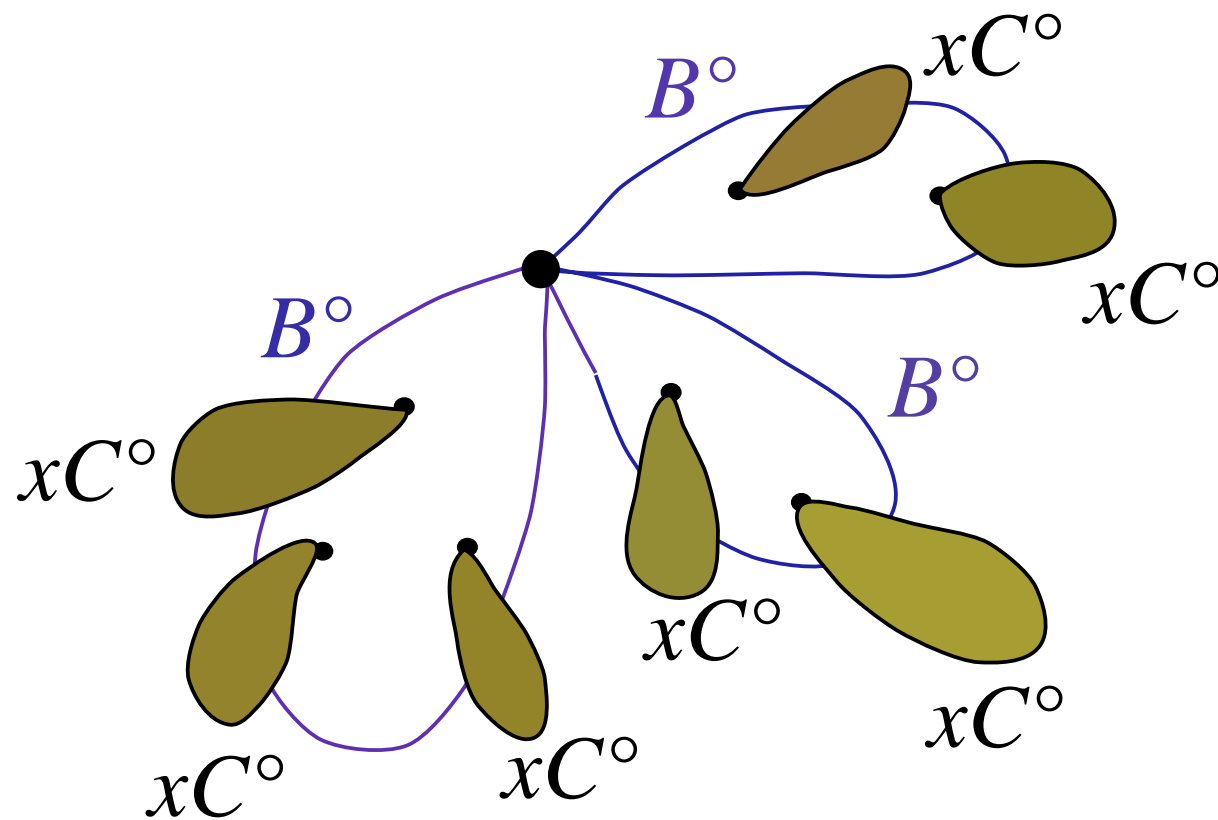
$$b_n = \frac{1}{2} a_{n-2} \cdot (n-1)! \quad (n \geq 3)$$

2-Connected Outer Planar Graphs



$$b_n = \frac{1}{2} a_{n-2} \cdot (n-1)! \quad (n \geq 3)$$

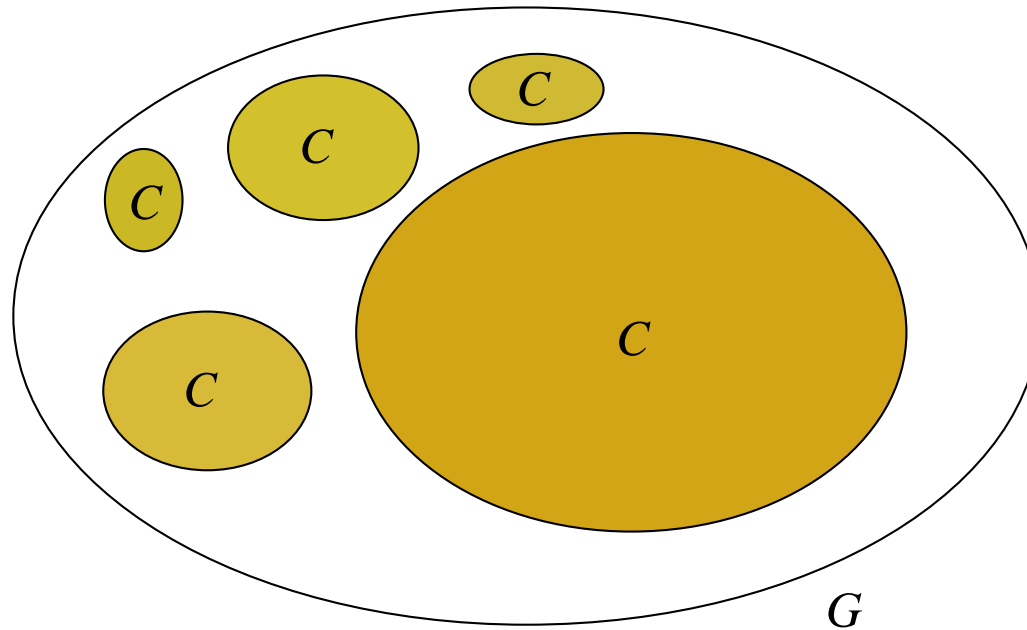
Connected Outer Planar Graphs



$$C^\bullet(x) = e^{B^\bullet(xC^\bullet(x))}$$

Connected Outer Planar Graphs

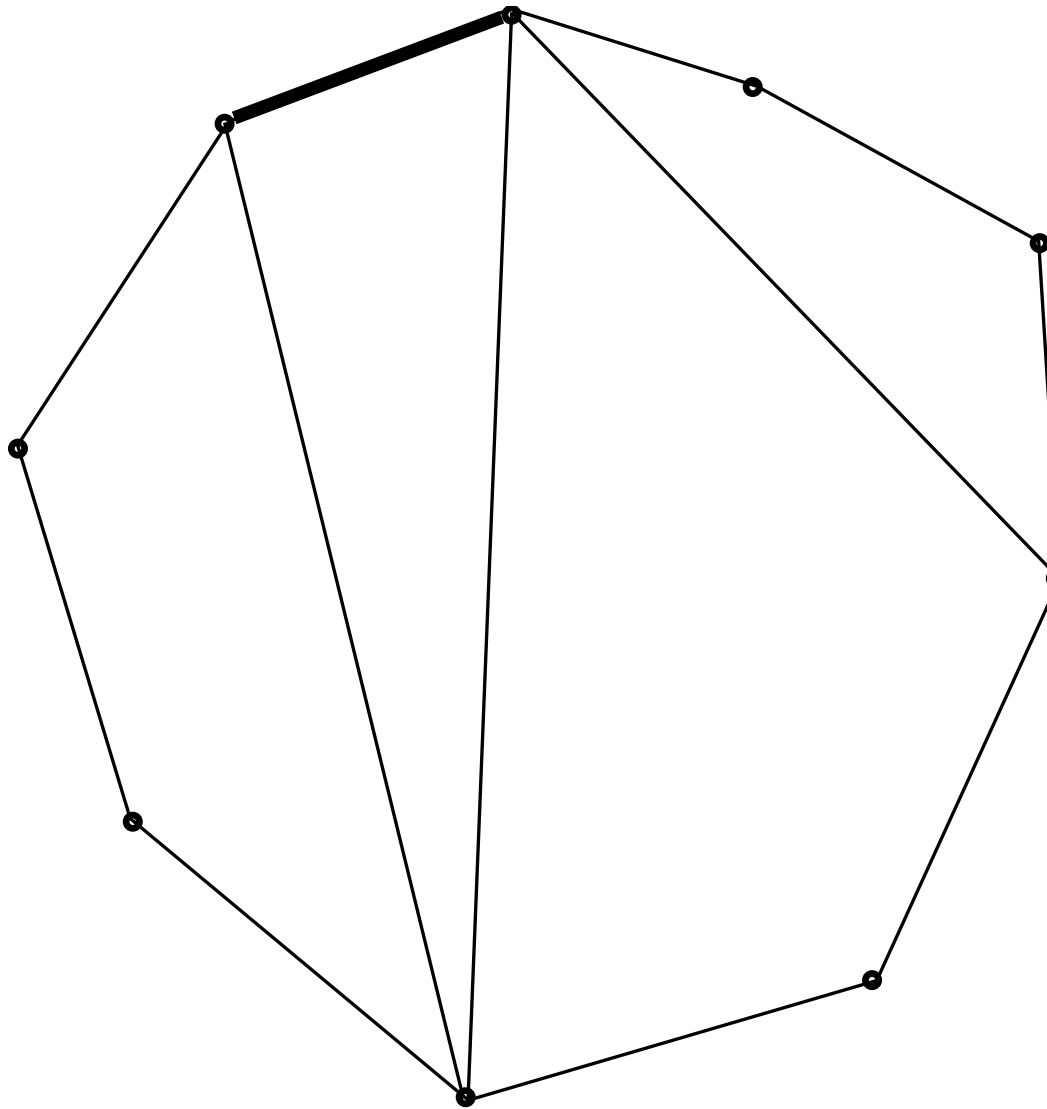
Generating Functions



$$G(x) = \exp(C(x))$$

Nodes of Given Degree

Dissections:



Nodes of Given Degree

- v_2 counts the number of nodes with degree 2,
- v_3 counts the number of nodes with degree 3,
- v counts the number of nodes with degree > 3 , and
- in all cases the two **nodes of the rooted edge** are **not taken into account**.

Nodes of Given Degree

- $A_{22}(v_2, v_3, v)$... generating function of dissections with the properties that both nodes of the rooted edge have degree 2,
- $A_{23}(v_2, v_3, v)$... generating function of dissections with the properties that the left node of the rooted edge has degree 2 and right one has degree 3,
- $A_{33}(v_2, v_3, v)$... generating function of dissections with the properties that both nodes of the rooted edge have degree 3,

Nodes of Given Degree

- $A_{2>}(v_2, v_3, v) \dots$ generating function of dissections with the properties that the left node of the rooted edge has degree 2 and the right has degree > 3 ,
- $A_{3>}(v_2, v_3, v) \dots$ generating function of dissections with the properties that the left node of the rooted edge has degree 3 and the right one has degree > 3 , and
- $A_{>>}(v_2, v_3, v) \dots$ generating function of dissections with the properties that both nodes of the rooted edge have degree > 3 .

Nodes of Given Degree

The sum

$$A(v_2, v_3, v) = A_{22} + 2A_{23} + A_{33} + 2A_{2>} + 2A_{3>} + A_{>>}$$

is the generating function of all dissections (with the corresponding counting in v_2, v_3, v).

In particular,

$$A(x) = A(x, x, x) = \frac{1 - 3x - \sqrt{1 - 6x + x^2}}{4x}.$$

Nodes of Given Degree

Lemma 3

$$A_{22} = v_2 + v_2 A_{22} + v_3 A_{23} + v A_{2>},$$

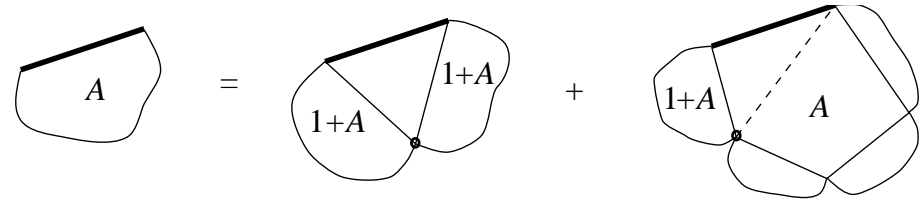
$$A_{23} = v_3 A_{22} + v(A_{23} + A_{2>}) = v_2 A_{23} + v_3 A_{33} + v A_{3>},$$

$$A_{33} = v(A_{22} + A_{23} + A_{2>})^2 + v(A_{22} + A_{23} + A_{2>})(A_{23} + A_{33} + A_{3>}),$$

$$A_{2>} = v_3 A_{23} + v(A_{33} + A_{3>}) + v(A_{2>} + A_{3>} + A_{>>}) + v_2 A_{2>} + v_3 A_{3>} + v A_{>>},$$

$$A_{3>} = v(A_{23} + A_{33} + A_{3>})(A_{2>} + A_{3>} + A_{>>}) + v(A_{22} + A_{23} + A_{2>})(A_{2>} + A_{3>} + A_{>>}),$$

$$A_{>>} = v(A_{23} + A_{33} + A_{3>} + A_{2>} + A_{3>} + A_{>>})^2 + v(A_{23} + A_{33} + A_{3>} + A_{2>} + A_{3>} + A_{>>})(A_{2>} + A_{3>} + A_{>>}).$$



Nodes of Given Degree

- $B_1^\bullet(v_1, v_2, v_3, v) \dots$ exponential generating functions of 2-connected labelled and marked outer planar graphs, where the root node (that is not counted) has degree 1.
- $B_2^\bullet(v_1, v_2, v_3, v) \dots$ exponential generating functions of 2-connected labelled and marked outer planar graphs, where the root node (that is not counted) has degree 2.
- $B_3^\bullet(v_1, v_2, v_3, v) \dots$ exponential generating functions of 2-connected labelled and marked outer planar graphs, where the root node (that is not counted) has degree 3.
- $B_{>3}^\bullet(v_1, v_2, v_3, v) \dots$ exponential generating functions of 2-connected labelled and marked outer planar graphs, where the root node (that is not counted) has degree > 3 .

Nodes of Given Degree

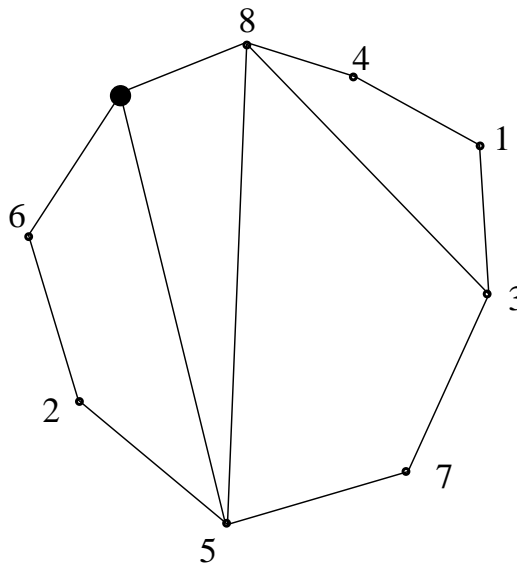
Lemma 4

$$B_1^\bullet(v_1, v_2, v_3, v) = v_1,$$

$$B_2^\bullet(v_1, v_2, v_3, v) = \frac{1}{2} (v_2 A_{22} + v_3 A_{23} + v A_{2>}),$$

$$B_3^\bullet(v_1, v_2, v_3, v) = \frac{1}{2} (v_2 A_{23} + v_3 A_{33} + v A_{3>}),$$

$$B_{>}^\bullet(v_1, v_2, v_3, v) = \frac{1}{2} (v_2 A_{2>} + v_3 A_{3>} + v A_{>>}).$$



Nodes of Given Degree

- $C_0^\bullet(v_1, v_2, v_3, v) \dots$ exponential generating functions of connected labelled and marked outer planar graphs, where the root node (that is not counted) has degree 0.
- $C_1^\bullet(v_1, v_2, v_3, v) \dots$ exponential generating functions of connected labelled and marked outer planar graphs, where the root node (that is not counted) has degree 1.
- $C_2^\bullet(v_1, v_2, v_3, v) \dots$ exponential generating functions of connected labelled and marked outer planar graphs, where the root node (that is not counted) has degree 2.

Nodes of Given Degree

- $C_3^\bullet(v_1, v_2, v_3, v) \dots$ exponential generating functions of connected labelled and marked outer planar graphs, where the root node (that is not counted) has degree 3.
- $C_{>3}^\bullet(v_1, v_2, v_3, v) \dots$ exponential generating functions of connected labelled and marked outer planar graphs, where the root node (that is not counted) has degree > 3 .

Nodes of Given Degree

Lemma 5

$$C_0^\bullet(v_1, v_2, v_3, v) = 1,$$

$$C_1^\bullet(v_1, v_2, v_3, v) = B_1^\bullet(W_1, W_2, W_3, W),$$

$$C_2^\bullet(v_1, v_2, v_3, v) = \frac{1}{2!} (B_1^\bullet(W_1, W_2, W_3, W))^2 + B_2^\bullet(W_1, W_2, W_3, W),$$

$$C_3^\bullet(v_1, v_2, v_3, v) = \frac{1}{3!} (B_1^\bullet(W_1, W_2, W_3, W))^3 \\ + \frac{1}{1!1!} B_1^\bullet(W_1, W_2, W_3, W) B_2^\bullet(W_1, W_2, W_3, W) \\ + B_3^\bullet(W_1, W_2, W_3, W),$$

$$C_{>}^\bullet(v_1, v_2, v_3, v) = e^{B_1^\bullet(W_1, W_2, W_3, W) + B_2^\bullet(\dots) + B_3^\bullet(\dots) + B_{>}^\bullet(W_1, W_2, W_3, W)} \\ - 1 - B_1^\bullet(W_1, W_2, W_3, W) - B_2^\bullet(\dots) - B_3^\bullet(\dots) \\ - \frac{1}{1!} (B_1^\bullet(W_1, W_2, W_3, W))^2 - \frac{1}{3!} (B_1^\bullet(W_1, W_2, W_3, W))^3 \\ - \frac{1}{1!1!} B_1^\bullet(W_1, W_2, W_3, W) B_2^\bullet(W_1, W_2, W_3, W),$$

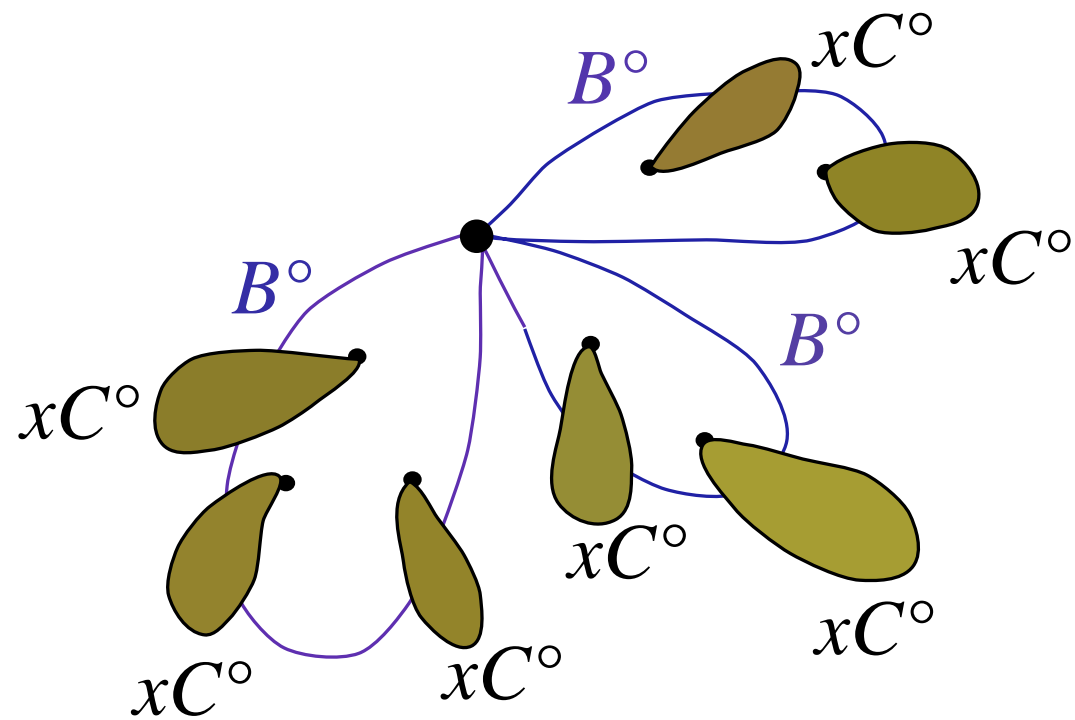
where on the right hand side

$$W_1 = v_1 C_0^\bullet + v_2 C_1^\bullet + v_3 C_2^\bullet + v(C_3^\bullet + C_>^\bullet),$$

$$W_2 = v_2 C_0^\bullet + v_3 C_1^\bullet + v(C_2^\bullet + C_3^\bullet + C_>^\bullet),$$

$$W_3 = v_3 C_0^\bullet + v(C_1^\bullet + C_2^\bullet + C_3^\bullet + C_>^\bullet),$$

$$W = v(C_0^\bullet + C_1^\bullet + C_2^\bullet + C_3^\bullet + C_>^\bullet).$$



Nodes of Given Degree

Counting nodes of degree 3:

$C(v_1, v_2, v_3, v)$... exponential generating function of all connected labelled outer planar graphs

$C_{d=3}(x, u)$... exponential generating function that counts the number of nodes with x and the number of nodes of degree $d = 3$ with u :

$$C_{d=3}(x, u) = C(x, x, xu, x).$$

Also:

$$\frac{\partial C_{d=3}(x, u)}{\partial x} = C_1^\bullet + C_2^\bullet + uC_3^\bullet + C_{>}^\bullet \quad \text{and} \quad \frac{\partial C_{d=3}(x, u)}{\partial u} = xC_3^\bullet$$

Thanks for your attention!