

TRAVELLING WAVES AND THE DISTRIBUTION OF THE HEIGHT OF BINARY SEARCH TREES*

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Outline of the Talk

- **Random Bisection Problem**
- **Binary Search Trees**
- **Results**
- **Travelling Wave**
- **Intersection Property**

Random Bisection Problem



Random Bisection Problem



Random Bisection Problem



Random Bisection Problem

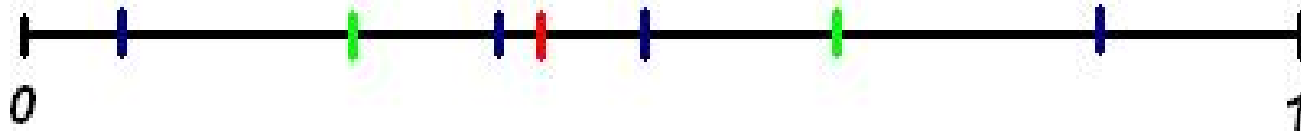


V is a random variable on $[0, 1]$

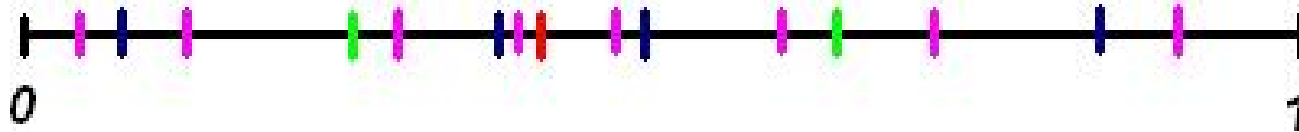
Random Bisection Problem



Random Bisection Problem



Random Bisection Problem



Random Bisection Problem

Analytic Description:

- x initial length (it was 1 in our example)
- $P_k(x, \ell)$ the probability that after k steps each of the 2^k fragments after is shorter (or equal) than ℓ
- $P_k(x, \ell) = P_k(x/\ell, 1) = \bar{P}_k(x/\ell)$
- $\bar{P}_0(x) = 1$ for $0 \leq x < 1$, $\bar{P}_0(x) = 0$ for $x \geq 1$

$$\bar{P}_{k+1}(x) = \mathbf{E} \left(\bar{P}_k(xV) \bar{P}_k(x(1-V)) \right)$$

Random Bisection Problem

t -Beta-Distribution for V :

$$f(x) = \frac{(2t + 1)!}{(t!)^2} \boxed{x^t (1 - x)^t}$$

density of distribution of V ($t \geq 0$ integer parameter)

Random Bisection Problem

Transfer Operator T :

$$(TF)(x) := E(F(xV)F(x(1-V)))$$

With this definition we have

$$\bar{P}_{k+1} = T\bar{P}_k$$

with $\bar{P}_0(x) = 1$ for $0 \leq x < 1$ and $\bar{P}_0(x) = 0$ for $x \geq 1$

Random Bisection Problem

Laplace Transform:

Set

$$L_F(u) := \int_0^\infty e^{-xu} F(x) dx \quad (u > 0)$$

Then

$$L_{\mathbf{T}F}(u)^{(2t+1)} = -\frac{(2t+1)!}{(t!)^2} \left(L_F(u)^{(t)}\right)^2$$

Random Bisection Problem

Laplace Transform:

Set

$$L_F(u) := \int_0^\infty e^{-xu} F(x) dx \quad (u > 0)$$

Then

$$L_{\mathbf{T}F}(u)^{(2t+1)} = -\frac{(2t+1)!}{(t!)^2} \left(L_F(u)^{(t)} \right)^2$$

and consequently

$$L_{\bar{P}_{k+1}}(u)^{(2t+1)} = -\frac{(2t+1)!}{(t!)^2} \left(L_{\bar{P}_k}(u)^{(t)} \right)^2$$

Binary Search Trees

Storing Data:

4, 6, 3, 5, 1, 8, 2, 7

Binary Search Trees

Storing Data:

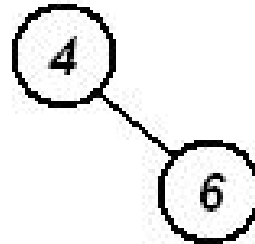
6, 3, 5, 1, 8, 2, 7

4

Binary Search Trees

Storing Data:

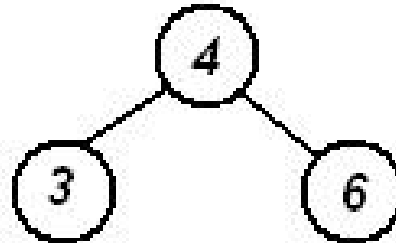
3, 5, 1, 8, 2, 7



Binary Search Trees

Storing Data:

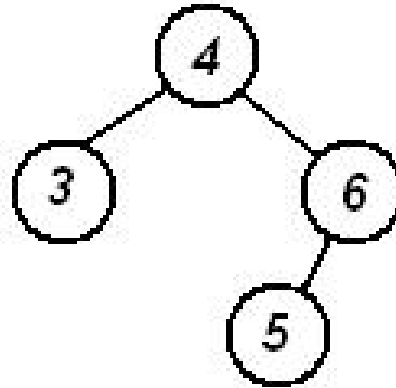
5, 1, 8, 2, 7



Binary Search Trees

Storing Data:

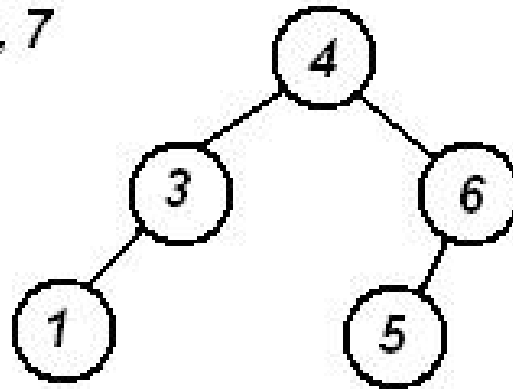
1, 8, 2, 7



Binary Search Trees

Storing Data:

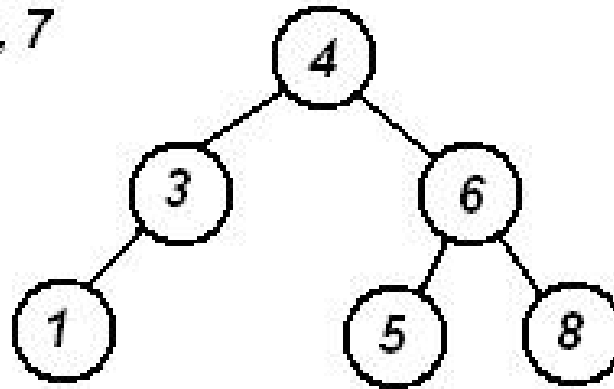
8, 2, 7



Binary Search Trees

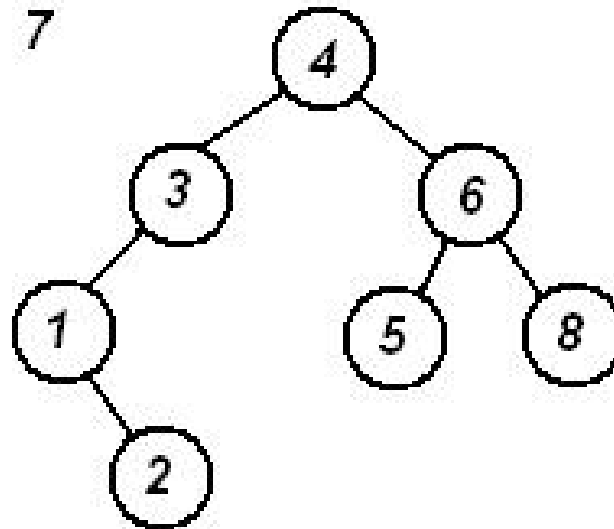
Storing Data:

2.7



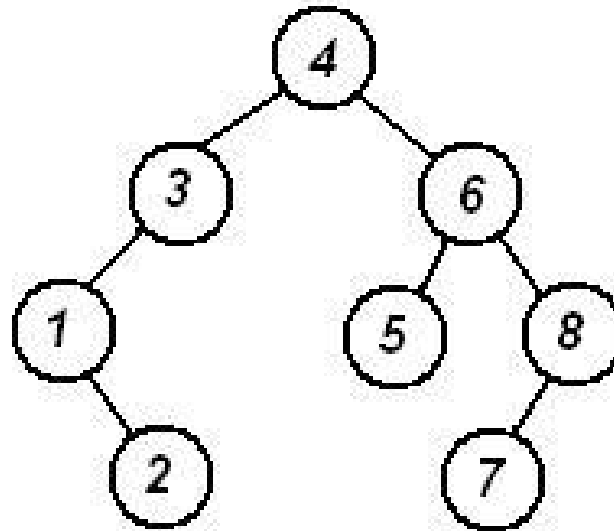
Binary Search Trees

Storing Data:



Binary Search Trees

Storing Data:



Binary Search Trees

Quicksort:

4, 6, 3, 5, 1, 8, 2, 7

Binary Search Trees

Quicksort:

4, 6, 3, 5, 1, 8, 2, 7

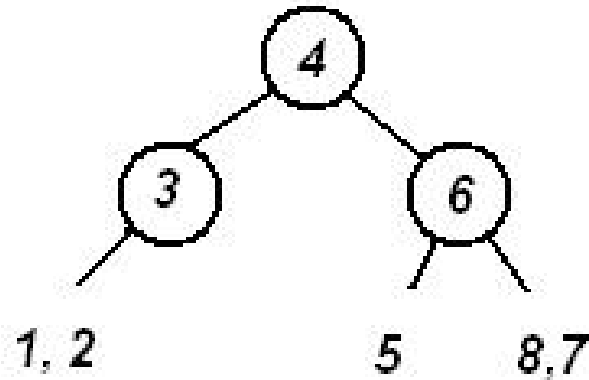


3, 1, 2

6, 5, 8, 7

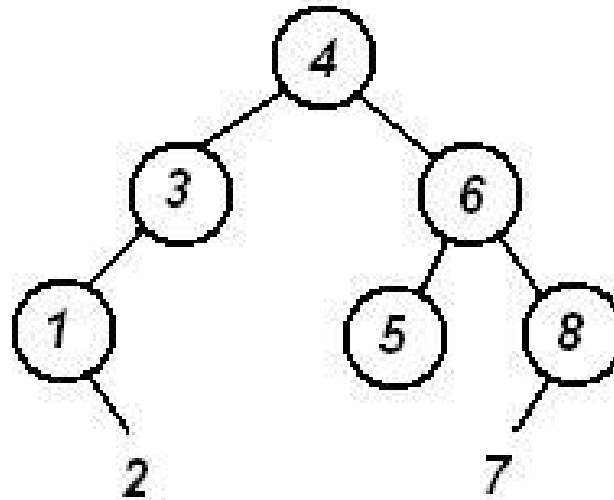
Binary Search Trees

Quicksort:



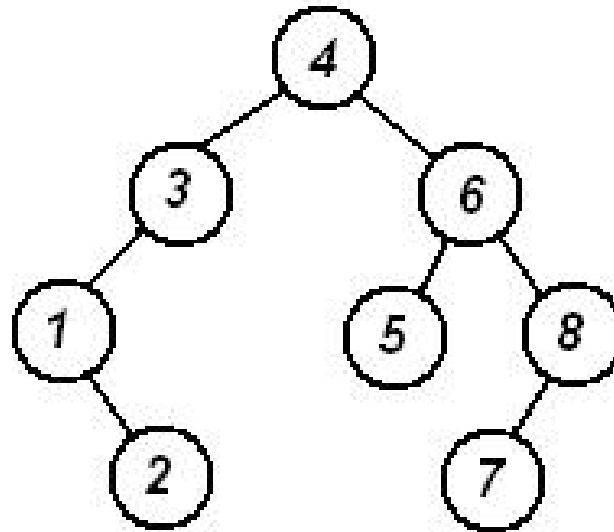
Binary Search Trees

Quicksort:



Binary Search Trees

Quicksort:



Binary Search Trees

Median of 3 – Quicksort:

4, 6, 3, 5, 1, 8, 2, 7

Binary Search Trees

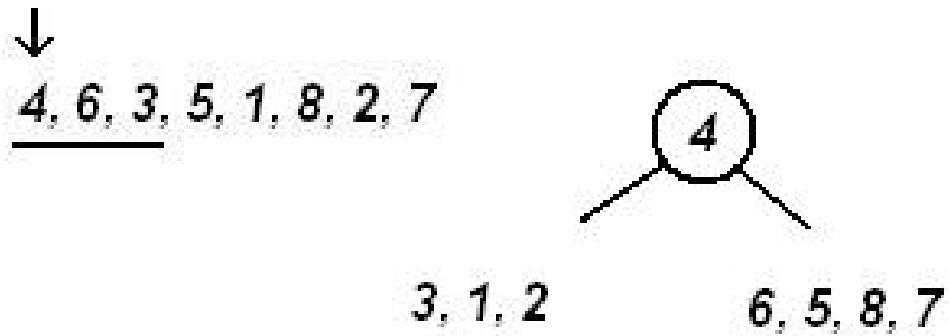
Median of 3 – Quicksort:



4, 6, 3, 5, 1, 8, 2, 7

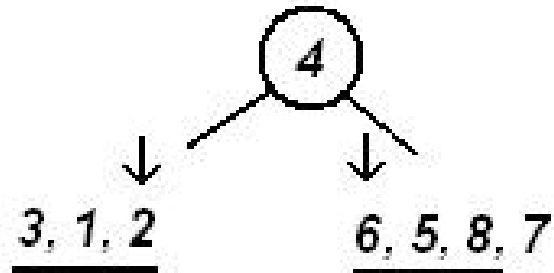
Binary Search Trees

Median of 3 – Quicksort:



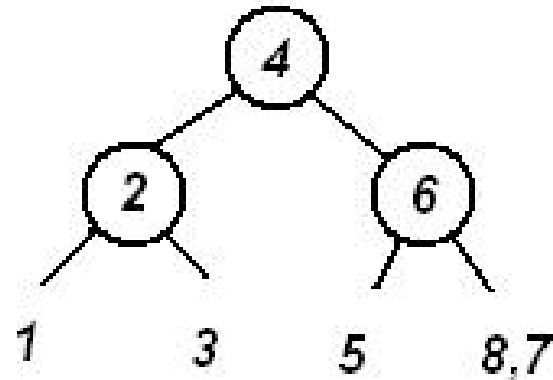
Binary Search Trees

Median of 3 – Quicksort:



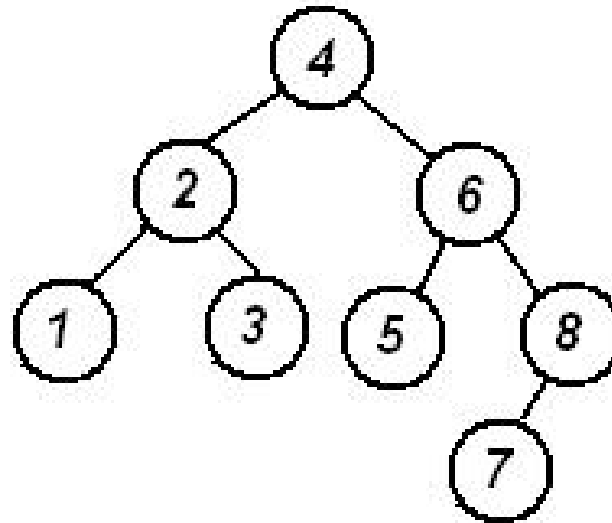
Binary Search Trees

Median of 3 – Quicksort:



Binary Search Trees

Median of 3 – Quicksort:



Binary Search Trees

Probabilistic Model:

Every permutation of $\{1, 2, \dots, n\}$ is equally likely.

→ probability distribution on binary trees of size n

→ every parameter on trees is a **random variable**

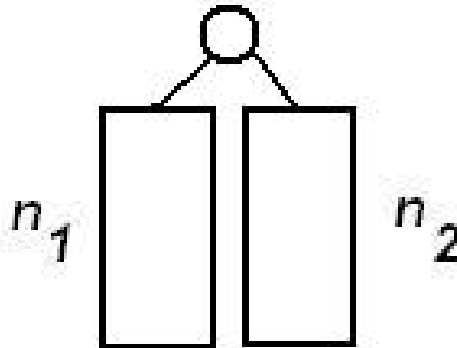
Notation

H_n ... **height** of trees (of size n)

Binary Search Trees

Height of Median of $(2t + 1)$ -Quicksort:
(fringe balanced binary search trees)

$$\Pr\{H_n \leq k + 1\} = \sum_{n_1+n_2=n-1} \frac{\binom{n_1}{t} \binom{n_2}{t}}{\binom{n}{2t+1}} \Pr\{H_{n_1} \leq k\} \cdot \Pr\{H_{n_2} \leq k\}$$



Binary Search Trees

Generating Functions:

$$y_k(x) = \sum_{n \geq 0} \Pr\{H_n \leq k\} \cdot x^n$$

$$y_{k+1}(x)^{(2t+1)} = \frac{(2t+1)!}{(t!)^2} \left(y_k(x)^{(t)}\right)^2$$

with initial conditions $y_0(x) = 1$, $y_k(0) = 1$.

Random Bisection Problem versus Binary Search Trees

Comparison:

Random bisection problem

$$L_{\bar{P}_{k+1}}(u)^{(2t+1)} = -\frac{(2t+1)!}{(t!)^2} \left(L_{\bar{P}_k}(u)^{(t)} \right)^2$$

Height of (fringe balanced) binary search trees

$$y_{k+1}(x)^{(2t+1)} = \frac{(2t+1)!}{(t!)^2} \left(y_k(x)^{(t)} \right)^2$$

Results

Theorem 1

Let x_k be defined by $\bar{P}_k(x_k) = \frac{1}{2}$. Then there exists a continuous function $F(x)$ such that (uniformly for $x \geq 0$ as $k \rightarrow \infty$)

$$\bar{P}_k(x) = F(x/x_k) + o(1).$$

More precisely, we have

$$x_k = e^{\rho k + \Theta(\log k)}$$

for some $\rho > 0$ (defined on the next slide) and $F(x)$ is uniquely defined by $F(1) = \frac{1}{2}$ and by the relation

$$F(x/\rho) = (\mathbf{T}F)(x).$$

Results

Definition of ρ :

Let $\beta > 0$ be the solution of

$$\sum_{j=0}^t \log(\beta + t + 1 + j) - \log(2t!) = \sum_{j=0}^t \frac{\beta}{\beta + t + 1 + j}.$$

Then

$$\rho = \sum_{j=0}^t \frac{1}{\beta + t + 1 + j}.$$

Results

Theorem 2

We have (uniformly for $k \geq 0$ as $n \rightarrow \infty$)

$$\Pr\{H_n \leq k\} = F(n/c_k) + o(1),$$

where c_k satisfies $c_k \sim cy_k(1)$ (for some $c > 0$) and

$$c_k = e^{\rho k + o(k)}$$

Furthermore,

$$\mathbf{E} H_n = \max\{k \geq 0 : c_k \leq n\} + O(1) \sim \frac{1}{\log \rho} \cdot \log n$$

and

$$\Pr\{|H_n - \mathbf{E} H_n| > y\} = O(e^{-\eta y}).$$

In particular we have, as $n \rightarrow \infty$, $\mathbf{Var} H_n = O(1)$.

Travelling Wave

First Observation:

If $F(x)$ satisfies

$$F(x/\rho) = (\mathbf{T}F)(x).$$

(for some $\rho > 0$) then

$$F_k(x) := F(x/\rho^k)$$

satisfies the recurrence

$$F_{k+1}(x) = (\mathbf{T}F_k)(x).$$

However, $F_0(x) = F(x) \neq \bar{P}_0(x) = \mathbf{1}_{[0,1]}(x)$

Travelling Wave

Second Observation:

Set

$$\Phi(u) = \int_0^{\infty} F(x)e^{-xu} dx$$

and

$$\tilde{y}_k(x) = e^{\rho k} \Phi(e^{\rho k}(1-x)).$$

Then

$$\tilde{y}_{k+1}(x)^{(2t+1)} = \frac{(2t+1)!}{(t!)^2} (\tilde{y}_k(x)^{(t)})^2.$$

Travelling Wave

Solution of $F(x/\rho) = (\mathbf{T}F)(x)$:

Set $A_1 = (e^{\rho V})^\beta$ and $A_2 = (e^{\rho(1-V)})^\beta$ and suppose that $X \geq 0$ satisfies the **stochastic fixed point equation**:

$$Y \stackrel{d}{=} A_1 Y_1 + A_2 Y_2$$

(where Y_1 and Y_2 have the same distribution as Y and $Y_1, Y_2, (A_1, A_2)$ are independent).

Then

$$F(x) = \mathbf{E} e^{-x^\beta Y}$$

satisfies

$$F(x/\rho) = (\mathbf{T}F)(x).$$

Travelling Wave

Proposition (Biggins, Kyprianou, Durrett, Liggett, ...)

Set

$$v(\alpha) = \log \left(\mathbf{E} \left(\sum_{i \geq 1} A_i^\alpha \right) \right)$$

and suppose that $v(0) > 0$, that $\alpha = 1$ is contained in the interior of $\{\alpha : v(\alpha) < \infty\}$, and that

$$v(1) = v'(1) = 0.$$

Then the stochastic fixed point equation

$$Y \stackrel{d}{=} \sum_{i \geq 1} A_i Y_i$$

has (up to scaling) a unique non-negative solution and the Laplace transform $\Phi(x) = \mathbf{E} e^{xY}$ satisfies

$$\lim_{x \rightarrow 0+} \frac{1 - \Phi(x)}{-x \log x} = c_1$$

for some constant $c_1 > 0$.

Travelling Wave

Remark 1.

The condition $v(1) = v'(1) = 0$ constitutes a **critical case**.

If $v(1) = 0$ and $v'(1) < 0$ then there is also a solution (that can be also obtained quite easily by a contraction argument).

If $v(1) = 0$ and $v'(1) > 0$ then there is **no solution**.

Remark 2.

The condition for β is just a reformulation that we are in the critical case.

Intersection Property

Point process:

$$Z = \sum_{j=1}^N \delta_{X_j},$$

Example: $N = 2$, $X_1 = \log(1/V)$, $X_2 = \log(1/(1 - V))$.

Transform \mathbf{T} (for distributions functions):

$$(\mathbf{T}G)(x) = \mathbf{E} \left(\prod_{j=1}^N G(x - X_j) \right).$$

Example: $G(x) = F(e^{-x})$: $F(x) = \mathbf{E}(F(xV)F(x(1 - V)))$.

Intersection Property

Intersection property:

*Suppose that $F(x)$ and $G(x)$ are continuous distribution functions such that the difference $F(x) - G(x)$ has exactly **one zero**. Then the difference $(\mathbf{T} F)(x) - (\mathbf{T} G)(x)$ has at most **one zero**.*

Intersection Property

Lemma.

Suppose that V is t -beta distributed and \mathbf{T} is defined by $(\mathbf{T}F)(x) = \mathbf{E}(F(xV)F(x(1 - V)))$.

Then the Laplace transforms $\Phi(u) = \int_0^\infty F(x)e^{-xu} dx$ satisfy an *intersection property*.

This property is the **key property** for the proof of Theorems 1 and 2.

It is not clear whether this is also true on the level of distributions functions?

Intersection Property

Theorem 3

Let $G_0(x) = 0$ for $x < 0$ and $G_0(x) = 1$ for $x \geq 0$ and set $G_{k+1} = \mathbf{T} G_k$, that is,

$$G_{k+1}(x) = \mathbf{E} \left(\prod_{j=1}^N G_k(x - X_j) \right).$$

If \mathbf{T} satisfies the *intersection property* then there exists $w(x)$ such that (uniformly for real x as $k \rightarrow \infty$)

$$G_k(x) = w(x - m(k)) + o(1),$$

where $m(k)$ is defined by $G_k(m(k)) = \frac{1}{2}$.

More precisely, we have

$$m(k) = kc + o(k).$$

for some constant $c > 0$ and $w(x)$ satisfies

$$w(x) = \mathbf{E} \left(\prod_{j=1}^N w(x + c - X_j) \right).$$

Thank You!